Strategic Supply Function Competition with Private Information

Xavier Vives*
IESE Business School
May 2011

Abstract
A finite number of sellers (n) compete in schedules to supply an elastic demand. The cost of each seller is random, with common and private value components, and the seller receives a private signal about it. A Bayesian supply function equilibrium is characterized: The equilibrium is privately revealing and the incentives to rely on private signals are preserved. Supply functions are steeper with higher correlation among the costs parameters. For high (positive) correlation supply functions are downward sloping, price is above the Cournot level, and as we approach the common value case price tends to the collusive level. As correlation becomes maximally negative we approach the competitive outcome. With positive correlation, private information coupled with strategic behavior induces additional distortionary market power above full information levels. Efficiency can be restored with appropriate subsidy schemes or with a precise enough public signal about the common value component. As the market grows large with the number of sellers the equilibrium becomes price-taking, bid shading is of the order of $1/n$, and the order of magnitude of welfare losses is $1/n^2$. The results extend to inelastic demand, demand uncertainty, and demand schedule competition. A range of applications in product and financial markets are presented.

Keywords: reverse auction, demand schedule competition, double auction, market power, adverse selection, competitiveness, public information, rational expectations, collusion, welfare.

JEL codes: L13, D44, D82, G14, L94, E58, F13

* I am very grateful to the editor and four referees for their extensive comments and feedback. I thank Carolina Manzano, Natalia Fabra, Juanjo Ganuza, John Moore, Jean Charles Rochet, Marzena Rostek and participants at seminars at Columbia, Harvard-MIT theory seminar, LSE, Yale, Toulouse, Chicago, Northwestern, Princeton, and ESSET at Gerzensee for very helpful comments, Rodrigo Escudero, Jorge Paz and Vahe Sahakyan for excellent research assistance, and the Department of Economics at MIT for its hospitality. The research in this paper has received funding from the European Research Council (European Advanced Grants scheme) under the European Community's Seventh Framework Programme (FP7/2007-2013) / ERC grant agreement n° 230254. Complementary support from Project Consolider-Ingenio CSD2006-00016 and project ECO2008-05155 of the Spanish Ministry of Science and Innovation at the Public-Private Sector Research Center at IESE, as well as the Barcelona GSE Research Network and the Generalitat de Catalunya is acknowledged.
1. Introduction

Many markets are characterized by traders competing in demand or supply schedules. This type of competition is very common in financial markets and some goods markets like wholesale electricity. Competition in supply functions has been used also to model bidding for government procurement contracts, management consulting, or airline pricing reservation systems, and provides a reduced form for strategic agency and trade policy models. Furthermore, the jury is still out on whether the price or quantity competition model fits better different oligopolistic markets and the supply function model appears as an attractive contender. In many of the situations depicted private information is relevant, uncertainty has both common and private values components, and traders are potentially strategic. This is the case, for example, in dealer markets; in both Treasury and central bank liquidity auctions, as well as in the reverse auctions proposed by the US Treasury to extract toxic assets from the balance sheets of banks. Supply function models in the industrial organization tradition have ignored private information while in the finance tradition demand function models have relied on the presence of noise traders.

In this paper I present a tractable model of strategic competition in schedules with an information structure which encompasses private and common values, avoiding the need to introduce noise traders or noisy supply, as well as paradoxes associated to fully revealing equilibria, and allowing a full welfare analysis. A main result is that private information with positive correlation of values generates market power, over and above the full information level, and this has deleterious welfare consequences. As a byproduct of the analysis we are able to explain market anomalies as well as provide policy prescriptions.

Consider a market where $n$ sellers compete in supply functions to satisfy a downward sloping demand. The market price is the one that equates aggregate supply and demand. Each seller receives a noisy signal of the uncertain intercept of his private marginal cost,
which has a non-random slope.¹ The modeling strategy is to consider linear-quadratic payoffs coupled with an affine information structure, which admits common and private value components, that yields a unique symmetric linear Bayesian supply function equilibrium (SFE) of the game among the \( n \) sellers. Linear equilibria are tractable, in particular in the presence of private information, have desirable properties like simplicity, and have proved very useful as a basis for empirical analysis. The characterization of a linear equilibrium with supply function competition when there is market power and private information needs some careful analysis in order to model the capacity of a seller to influence the market price at the same time that the seller learns from the price. Kyle (1989) pioneered this type of analysis in a financial market context introducing noise trading in order to prevent prices from being fully revealing and the market collapsing. The present paper provides a tractable alternative to the models with an aggregate exogenous shock and is based on rational traders who are heterogeneous because of idiosyncrasies that translate in their market positions.

It is found that there is a unique SFE except in the limit cases of pure common value and maximal negative correlation. This equilibrium is privately revealing. That is, the private information of a firm and the price provide a sufficient statistic of the joint information in the market. This means in particular that each trader has incentives to rely on his private signal despite the fact that the price aggregates information about the signals of other traders in the market.

In the linear equilibrium sellers, when they have private information on their positively correlated costs, are more cautious when they see a price increase since it may mean that costs are high. The more so, with sellers using steeper schedules, when signals are noisier or costs parameters more correlated. The market looks less competitive in those circumstances as reflected in increased expected price-cost margins. This is reminiscent of the winner’s curse in auctions. Indeed, the price has an information role on top of its

¹ We could consider as well the symmetric situation where \( n \) buyers with uncertain private valuations compete to fulfill an upward sloping supply schedule. In the paper we stick to the seller convention until we develop applications.
traditional capacity as index of scarcity. In fact, when the first effect is strong enough supply functions slope downwards and prices are above the Cournot level. This is in contrast to the results of Klemperer and Meyer (1989) with symmetric information. More surprisingly perhaps, as we approach the common value case the price tends to the collusive level. This is because of information-induced market power at the unique linear equilibrium and not because of the existence of a vast multiplicity of equilibria. Even with constant marginal costs there is market power in equilibrium when adverse selection is severe enough. Relaxation of competition due to adverse selection is also obtained in Biais et al. (2000) in a different pure common value environment. When costs are negatively correlated then there is “favorable” selection and competition is intensified with respect to the full information benchmark, the competitive outcome being attained with maximal negative correlation.

Sellers at the strategic equilibrium act as if there were price–takers but facing steeper marginal costs than the true ones. The distortion has a full information market power component and another component induced by private information, which is increasing with the correlation of cost shocks and noise in the signals (when costs are positively correlated). Both aggregate/allocative and distributive/productive inefficiency increase with the size of the distortion, implying too low sales and too similar sales among sellers. As we approach the common value case expected profits converge to the collusive level. Furthermore, simulations suggest that typically the expected deadweight loss increases as we approach the common value case and with noisier signals. A welfare optimal allocation can be implemented by a price-taking Bayesian supply function equilibrium. It is shown how a quadratic subsidy which lowers the perceived slope of marginal costs of sellers may induce price-taking behavior and restore efficiency.

The paper considers also the large market case where the number of sellers and demand are replicated (with \( n \) the number of sellers and the size of the market as well). Then both the distortion induced by private information and bid shading are decreasing in \( n \).
Furthermore, bid shading is of the order of $1/n$, in a large market there is no efficiency loss (in the limit), and the order of magnitude of the expected deadweight loss is $1/n^2$.\textsuperscript{2}

The welfare evaluation of the SFE is in marked contrast with the Cournot equilibrium in the presence of private information. The reason is that the SFE aggregates information but not the Cournot market.\textsuperscript{3} In a large Cournot market, in general, there is a welfare loss due to private information even in the limit.

The results are shown to be robust to a number of extensions: costly information acquisition, with maintained incentives under certain conditions even when close to the common value case; inelastic demand, which makes the model basically a double auction similar to Kyle (1989) but with no need of noise traders; demand uncertainty, which makes the equilibrium noisy and shows how an increase in noise in the public statistic lessens the adverse selection problem (when there is positive correlation); and the introduction of a public signal, which if it is precise enough may restore efficiency. The model with demand uncertainty has as a limit, with appropriate choice of parameters, the markets considered in the linear Klemperer and Meyer (1989) model and in the Kyle (1989) risk neutral informed traders model.

A leading application of the model to goods markets is to wholesale electricity. The model admits also other interpretations. The cost shock could be related to some ex post pollution or emissions damage which is assessed on the firm, or it could be a random opportunity cost of serving the market which is related to revenue management dynamic considerations. At the same time the reinterpretation of the results in terms of demand schedule competition opens up a host of applications to financial markets (e.g. to legacy

\begin{footnotesize}
\begin{itemize}
\item[2] This is also the rate of convergence to efficiency obtained in a double auction context by Cripps and Swinkels (2006). Vives (2011b) deals with the limit continuum economy case and provides a foundation for competitive rational expectations equilibria.
\item[3] The welfare analysis in the supply function model contrasts thus with the one in models where there is no endogenous public signal such as the Cournot market in Vives (1988), the beauty contest in Morris and Shin (2002), or the general linear-quadratic set up of Angeletos and Pavan (2007).
\end{itemize}
\end{footnotesize}
loans, central bank liquidity, and Treasury auctions). Each of these applications is dealt with in Section 6.

Competition in supply or demand schedules has a long tradition in the literature. It has been studied in the absence of uncertainty by Grossman (1981) and Hart (1985) showing a great multiplicity of equilibria.\(^4\) Similar results in a complete information setting are obtained by Wilson (1979) in a share auction model and by Bernheim and Whinston (1986) in a menu auction. Back and Zender (1993) and Kremer and Nyborg (2004) obtain related results for Treasury auctions. Some of the equilibria can be very collusive.\(^5\) Klemperer and Meyer (1989) show how adding uncertainty in the supply function model can reduce the range of equilibria and even pin down a unique equilibrium (linear in a linear-quadratic model) provided the uncertainty has unbounded support. In this case the supply function equilibrium is always between the Cournot and competitive (Bertrand) outcomes.\(^6\) The supply function models considered typically do not allow for private information.\(^7\) Kyle (1989) introduces private information into a double auction for a risky asset of unknown liquidation value and derives a unique symmetric linear Bayesian equilibrium in demand schedules when traders have constant absolute risk aversion, there is noise trading, and uncertainty follows a Gaussian distribution.

The plan of the paper is as follows. Section 2 presents the supply function model with strategic sellers and characterizes a SFE and its comparative static properties. Section 3 performs a welfare analysis characterizing the distortion at the SFE and deadweight losses, including welfare simulations and a comparison with Bayesian Cournot equilibria, and showing how the efficient allocation can be attained with price-taking equilibria and

---

\(^4\) Grossman thought of firms signing implicit contracts with consumers that committed the firm to a supply function. Hart uncovers the equivalence between choosing a reaction function and a supply function.

\(^5\) Back and Zender (2001) and LiCalzi and Pavan (2005) show how the auction can be designed to limit those collusive equilibria.

\(^6\) This is also the result in Vives (1986) where the slope of the supply function is fixed by technological considerations.

\(^7\) Exceptions are the empirical papers of Hortaçsu and Puller (2008) and Kühn and Machado (2004) in electricity.
implemented with subsidy schemes. Section 4 studies replica markets and characterizes the convergence to price-taking behavior as the market grows large and the order of magnitude of deadweight losses. Section 5 deals with the extensions: inelastic demand, demand uncertainty, public signals, and demand schedule competition. Section 6 develops the applications. Concluding remarks, including potential policy implications, close the paper. Proofs are gathered in the Appendix and in the Supplement (Vives (2011a)), which includes also details of the simulations of the model, the analysis of endogenous information acquisition and the Bayesian Cournot model.

2. A strategic supply function model

Consider a market for a homogenous good with $n$ sellers. Seller $i$ faces a cost

$$C(x_i; \theta_i) = \theta_i x_i + \lambda x_i^2$$

of supplying $x_i$ units of the good where $\theta_i$ is a random parameter and $\lambda > 0$. Demand arises from an aggregate buyer with quasilinear preferences and gross surplus $U(y) = \alpha y - \beta y^2 / 2$, where $\alpha$ and $\beta$ are positive parameters and $y$ the consumption level. This gives rise to the inverse demand $P(y) = \alpha - \beta y$. In a reverse auction, for example, the buyer presents the schedule $P(y) = \alpha - \beta y$ to the sellers who will bid to supply. Total surplus is therefore given by $TS = U(\sum_i x_i) - \sum_i C(x_i, \theta_i)$.

We assume that $\theta_i$ is normally distributed (with mean $\alpha > \bar{\theta} > 0$ and variance $\sigma_\theta^2$). The parameters $\theta_i$ and $\theta_j$, $j \neq i$, are correlated with $\text{cov}[\theta_i, \theta_j] = \rho \sigma_\theta^2$, $\rho \in [-\frac{1}{n-1}, 1]$, for $j \neq i$. The average parameter $\bar{\theta} \equiv \left(\sum_{i=1}^n \theta_i\right)/n$ is thus normally distributed with

---

8 We could also deal easily with the case where there the seller faces an adjustment cost of the form $\lambda (x_i - \hat{x}_i)^2 / 2$ where $\hat{x}_i$ is a target quantity for agent $i$.

9 We will comment in Section 5.1 on how the results specialize to the case of inelastic demand and in Section 5.5 on how they can be reinterpreted for the case of demand instead of supply bids.
mean $\bar{\theta}$, $\text{var} [\bar{\theta}] = (1 + (n - 1)\rho)\sigma_\theta^2$, and $\text{cov} [\bar{\theta}, \theta] = \text{var} [\bar{\theta}]$.\(^{10}\) Seller $i$ receives a signal $s_i = \theta_i + \varepsilon_i$ with $\varepsilon_i$ normally distributed, $E[\varepsilon_i] = 0$ and $\text{var} [\varepsilon_i] = \sigma_\varepsilon^2$. Error terms in the signals are uncorrelated among themselves and with the $\theta_i$ parameters. Ex-ante, before uncertainty is realized, all sellers face the same prospects.\(^{11}\)

Our information structure encompasses the cases of “common value” and of “private values”. For $\rho = 1$, the $\theta$ parameters are perfectly correlated and we are in a common value model. When signals are perfect, $\sigma_\varepsilon^2 = 0$ for all $i$, and $0 < \rho < 1$, we are in a private values model. Agents receive idiosyncratic shocks, which are imperfectly correlated, and each agent observes his shock with no measurement error. When $\rho = 0$, the parameters are independent, and we are in an independent values model. When $\rho < 0$, the costs parameters are negatively correlated. The case of non-negative correlation is the one more relevant empirically.

2.1 Equilibrium
Sellers compete in supply functions. We will restrict attention to symmetric linear Bayesian supply function equilibrium (SFE for short).\(^{12}\) The strategy for seller $i$ is a price contingent schedule $X(s_i, \cdot)$. This is a map from the signal space to the space of supply functions. Given the strategies of sellers $j = 1, \ldots, n$ for given realizations of signals market clearing implies that $p = P\left(\sum_{j=1}^{n} X(s_j, p)\right)$. Let us assume that there is a

\(^{10}\) Note that $\text{var} [\bar{\theta}] \geq 0$ for $\rho \geq -(n - 1)^{-1}$ since then $1 + (n - 1)\rho \geq 0$.

\(^{11}\) With normal distributions there is positive probability that prices and quantities are negative in equilibrium. This can be controlled by choice of the variances of the distributions and the parameters $\alpha , \beta , \lambda$ and $\bar{\theta}$.

\(^{12}\) What makes the model tractable is the combination of linear-quadratic payoffs coupled with an affine information structure (that is, a pair of prior and likelihood that yields affine conditional expectations as under the normality) that allows for the existence of linear equilibria. It is crucial that the slopes of demand and costs are not affected by uncertainty. Adding (intercept) demand uncertainty presents no problem as long as the affine information structure is kept (see Section 5.2).
unique market clearing price \( \hat{p}(X(s_1, \cdot), \ldots, X(s_n, \cdot)) \) for any realizations of the signals.\(^{13}\)

Then profits for seller \( i \), for any given realization of the signals, are given by

\[
\pi_i(X(s_1, \cdot), \ldots, X(s_n, \cdot)) = pX(s_i, p) - C(X(s_i, p))
\]

where \( p = \hat{p}(X(s_1, \cdot), \ldots, X(s_n, \cdot)) \). This defines a game in supply functions and we want to characterize a SFE. Given linear strategies of rivals \( X(s_j, p) = b - as_j + cp, \quad j \neq i \), seller \( i \) faces a residual inverse demand

\[
p = \alpha - \beta \sum_{j \neq i} X(s_j, p) - \beta x_i = \alpha - \beta (n-1)(b+cp) + \beta a \sum_{j \neq i} s_j - \beta x_i.
\]

Provided \( 1 + \beta(n-1)c > 0 \) it follows that \( p = I_i - dx_i \), where \( I_i = d \left( \alpha \beta^{-1} - (n-1)b + a \sum_{j \neq i} s_j \right) \), and \( d = \left( \beta^{-1} + (n-1)c \right)^{-1} \). The (endogenous) parameter \( d \) is the (absolute value of the) slope of inverse residual demand for a seller and plays an important role in the characterization of equilibrium and its welfare properties. All the information provided by the price to seller \( i \) about the signals of others is subsumed in the intercept of residual demand \( I_i \). The expression for residual demand disentangles the capacity of a seller to influence the market price (\( d \)) from learning from the price (\( I_i \)). Note that \( I_i \) is informationally equivalent to \( h_i = \beta a \sum_{j \neq i} s_j \).

The information available to seller \( i \) is therefore \( \{s_i, p\} \) or, equivalently, \( \{s_i, h_i\} \). Seller \( i \) chooses \( x_i \) to maximize

\[
E[\pi_i|s_i, p] = x_i \left( p - E[\theta|s_i, p] \right) - \frac{\lambda}{2} x_i^2 = x_i \left( I_i - dx_i - E[\theta|s_i, p] \right) - \frac{\lambda}{2} x_i^2.
\]

The F.O.C. is

\[
I_i - E[\theta|s_i, I_i] - 2dx_i - \lambda x_i = 0 \quad \text{or, equivalently,} \quad p - E[\theta|s_i, p] = (d + \lambda) x_i.
\]

The second order sufficient condition for a maximum is \( 2d + \lambda > 0 \). An equilibrium must fulfill also \( 1 + \beta(n-1)c > 0 \). The following proposition characterizes the linear equilibrium and the following subsections its properties.\(^{14}\)

\(^{13}\) If there is no market clearing price assume the market shuts down and if there are many then the one that maximizes volume is chosen.

\(^{14}\) We use the term increasing or decreasing in the strict sense unless otherwise stated.
Proposition 1. Let \(- (n-1)^{-1} < \rho < 1\) and \(\sigma^2 / \sigma^2_0 < \infty\).

(i) If \(\lambda > 0\) there is a unique SFE. It is given by the supply function \(X(s, p) = \left( p - E[\theta | s, p] \right) / (d + \lambda) \) with \(d = \left( \beta^{-1} + (n-1)c \right)^{-1}\), and \(c = \partial X / \partial p\) is given by the largest solution to a quadratic equation \(g(c; M) = 0\) where

\[
M = \frac{n \rho \sigma^2_e}{(1-\rho)(\sigma^2_e + (1+(n-1)\rho)\sigma^2_\theta)}. \quad \text{We have that} \quad a = -\partial X / \partial s_i > 0, \quad 0 < d < \beta n, \quad c > -M \left( (1+M) \beta n \right)^{-1}, \quad 1+M > 0; \quad c \text{ decreases with } \lambda \text{ and } M, \text{ ranging from } -(\beta n)^{-1} \text{ to } \infty \text{ as } M \text{ ranges from } \infty \text{ to } -1.
\]

(ii) If \(\lambda = 0\) there is a SFE if and only if \(n-M-2<0\), then \(c = c_0 = -\left( n-M \right) / \left( (n-M-2) \beta n \right) \).

2.2 Information revelation

The equilibrium price \(p\) is a linear function of, and therefore reveals, the aggregate information \(\bar{s} \equiv \left( \sum_i s_i \right) / n \).\(^{15}\) The equilibrium is *privately revealing*. That is, for seller \(i\), \((s_i, p)\) or \((s_i, \bar{s})\) is a sufficient statistic of the joint information in the market \(s = (s_1, \ldots, s_n)\) in the estimation of \(\theta_i\) (see Allen (1981)). In particular, in equilibrium we have that the conditional distribution of posterior beliefs of \(\theta_i\) fulfils \(E[\theta | s_i, p] = E[\theta | s_i, \bar{s}] = E[\theta | \bar{s}] \).\(^{16}\) That is, a seller obtains from the price the collective information of other sellers (which is relevant as long as costs are correlated) but still his private signal is useful to improve the estimation of his cost parameter (provided \(\rho < 1\)). This means that incentives to rely (and purchase) private signals

\(^{15}\) Average quantity is given by \(\bar{x} = \left( \sum_i x_i \right) / n = b - a \bar{s} + cp\). Substituting in the inverse demand \(p = \alpha - \beta n \bar{x}\), noting that in equilibrium \(1 + \beta nc > 0\), and solving for \(p\) we obtain \(p = (1 + \beta nc)^{-1} (\alpha - \beta nb + \beta na \bar{s})\).

\(^{16}\) Note that under normality the conditional expectation is a sufficient statistic for the conditional distribution.
remain since a private signal adds information for seller $i$ on top of the information conveyed by the price. (Indeed, we have that $a \equiv -\partial X/\partial s_i > 0$ if $\rho < 1$.)

If the signals are costly to acquire and agents face a convex cost of acquiring precision $\tau_\epsilon \equiv 1/\sigma_\epsilon^2$ then it is possible to show that each seller will have an incentive to purchase some precision for any $-(n-1)^{-1} < \rho < 1$ and any $n$ provided that the marginal cost of acquiring precision is small enough for little amounts of precision or that the prior is diffuse enough. The reason is as follows. A seller by purchasing a signal will improve the information on his random cost parameter even though he learns the signals of the other sellers through the price. When the number of sellers is large or correlation is high the improvement will be small but if the seller can purchase a little bit of precision at a small cost he will do it. Furthermore, the more diffuse is the prior the higher the marginal value of information. 17

An equivalent formulation that highlights the aggregate and idiosyncratic components of uncertainty is to let $\theta_i \equiv \theta_i - \tilde{\theta}$ and note that $\theta_i = \tilde{\theta} + \vartheta_i$, where $\text{cov} \left[ \vartheta_i, \tilde{\theta} \right] = 0$ and $n^{-1} \sum_{i=1}^{n} \vartheta_i = 0$. 18 It becomes clear then that key to the private revealing property of the equilibrium is that the same signal $s_i$ conveys information about the idiosyncratic component $\eta_i$ and an aggregate component $\tilde{\theta}$, and that the price reveals a sufficient statistic of the signals of sellers other than $i$. 19

---

17 If the marginal cost of acquiring precision is positive at zero and $\rho$ close to 1 then for $n$ large enough there is no purchase of information. However, this is not the case in the natural case of a large market where the number of buyers and sellers grow together (as in Section 4). See Section S.4 in the Supplement (Vives (2011a)) for the information acquisition model, results and proofs.

18 We could also let $\hat{\theta}_i \equiv \theta_i - \tilde{\theta}$, where $\tilde{\theta}_i \equiv \lim_{n \to \infty} n^{-1} \sum_{i=1}^{n} \theta_i$ is now the common component, then we would have both $\text{cov} \left[ \hat{\theta}_i, \tilde{\theta} \right] = 0$ and $\text{cov} \left[ \hat{\theta}_i, \hat{\theta}_j \right] = 0$ for $i \neq j$.

19 This latter property obtains typically in the linear-Gaussian models with uniform correlation in the parameters (see, e.g., Ch. 3 in Vives (2008)). However, Rostek and Weretka (2010) show that this need not be the case with heterogeneous correlation (when considering symmetric equilibria which depend only on average correlation).
The equilibrium is in contrast with the pure common value “noise trader” model of Kyle (1989) where risk-averse traders, some with private information about the value of the risky asset, face liquidity traders. Here the collective information of traders would be revealed by the price, and the market would collapse, except for the presence of noise traders. In our base model there is no shock to the residual demand function (be it noise traders or demand uncertainty) and, consequently, in the pure common value case ($\rho = 1$ and $\sigma^2_\varepsilon < \infty$) the equilibrium collapses.\(^{20}\) Indeed, when $\rho = 1$ and $0 < \sigma^2_\varepsilon < \infty$ a fully revealing rational expectations equilibrium is not implementable and there is no linear equilibrium. The reason should be well understood: If the price reveals the common value then no seller has an incentive to put any weight on his signal (and the incentives to acquire information disappear as well). But if sellers put no weight on their signals then the price can not contain any information on the costs parameters. This is the essence of the Grossman-Stiglitz paradox (1980). The approach in our paper allows performing a welfare analysis since it does away with the need to introduce noise traders who do not have a well-defined utility function. The private value component in the valuation of an agent in our model arises naturally in applications as we will see in Section 6.

2.3 Private information and market power

Despite the fact that the SFE is privately revealing it is distorted in relation to the full information supply function equilibrium where sellers share $s = (s_1, \ldots, s_n)$ (denoted by a superscript $f$). Indeed, following a similar analysis as before it is easy to see that $X^f(s, p) = \left(p - E[\theta_i|s_i]\right) / (d^f + \lambda)$ where $d^f$ and $c^f$ correspond in Proposition 1 to the case $M = 0$. Whenever there is no correlation between the cost parameters ($\rho = 0$) or signals are perfect (the private values case with $\sigma^2_\varepsilon / \sigma^2_\theta = 0$)\(^{21}\), $M = 0$.

\(^{20}\) As we will see in Section 2.4, the equilibrium also collapses when $\rho = -(n-1)^{-1}$ since then there is no aggregate uncertainty (var $\bar{\theta} = 0$).

\(^{21}\) In this case the equilibrium is independent of $\rho$ and it exists even if $\rho = 1$ or $\rho = -(n-1)^{-1}$. 

\[ E[\theta_i|s_i, p] = E[\theta_i|s] = E[\theta_i|s_i] \] and seller \( i \) does not learn about \( \theta_i \) from prices. In these cases the SFE coincides with the full information equilibrium.\(^{22}\)

When \( \rho \sigma^2_\varepsilon \neq 0 \), the price at a SFE serves a dual role as index of scarcity and as conveyor of information. This can be seen from the supply function \( X(s_i, p) = \left( p - E[\theta_i|s_i, p] \right) / (d + \lambda) \). Indeed, a high price has a direct effect to increase the competitive supply of a seller, but also conveys news that costs for the seller are high if \( \rho > 0 \) (since then \( E[\theta_i|s_i, p] \) is increasing in \( p \)) or low if \( \rho < 0 \) (since then \( E[\theta_i|s_i, p] \) is decreasing in \( p \)). When \( \rho \sigma^2_\varepsilon > 0 \) supply functions at the SFE are steeper than with full information \( c < c^f \) (and \( d > d^f \)) due to adverse selection. Private information creates market power \( (d) \) over and above the full information level \( (d^f) \).

When \( \rho \sigma^2_\varepsilon < 0 \) supply functions are flatter \( c < c^f \) and the informational effect of the price is pro-competitive since a high price conveys the news to a seller that the costs of rivals are high and therefore that his own costs are low due to negative correlation. As \( \rho \) increases from \( \rho = 0 \) the adverse selection problem worsens and as \( \rho \) turns negative the adverse selection problem disappears and becomes “favorable” selection. The parameter \( M \) (a function of \( \rho \) and \( \sigma^2_\varepsilon / \sigma^2_\theta \)) is an index of adverse selection and the slope of the supply function becomes steeper \( (c \) decreases) with \( M \) (Proposition 1). We have that when \( \sigma^2_\varepsilon > 0 \), \( M \) is increasing in \( \rho \) and \( \text{sgn}\left\{ \partial M / \partial \left( \sigma^2_\varepsilon / \sigma^2_\theta \right) \right\} = \text{sgn}\{ M \} = \text{sgn}\{ \rho \}. \)

Both as either \( |\rho| \) increases and cost parameters become more correlated, or \( \sigma^2_\varepsilon / \sigma^2_\theta \) increases and private signals are (relatively) less precise, the price signal becomes more relevant to estimate \( \theta_i \). More precisely, the absolute value of the weight on the information component of the price \( h_i \) in \( E[\theta_i|s_i, h_i] \) increases in \( |\rho| \) and \( \sigma^2_\varepsilon / \sigma^2_\theta \).\(^{23}\) When \( \rho \sigma^2_\varepsilon > 0 \) the

\(^{22}\) This equilibrium is robust to the introduction of noise in the demand function as in Klemperer and Meyer (1989), see Section 5.2.

\(^{23}\) See the proof of Claim A.1 in the Appendix.
result is that as $\rho$ or $\sigma^2_e/\sigma^2_\theta$ increase then $c$ decreases since a high price is bad news (i.e., the seller learns more from the price about its cost shock and reacts less to a price change than if the price was only an index of scarcity). When $\rho\sigma^2_e < 0$, as $|\rho|$ or $\sigma^2_e/\sigma^2_\theta$ increases then $c$ increases since a high price is good news.\(^{24}\)

As $\rho$ tends to 1, $M$ tends to $\infty$ and $c$ becomes negative. As $\rho \to 1$ we have that the equilibrium collapses in the limit. In fact, since $a \to 0$, $c \to -(\beta n)^{-1}$, $b \to \alpha (\beta n)^{-1}$, the supply function of a seller converges to the per capita seller demand function $x = (\alpha - p)(\beta n)^{-1}$ (see Figure 1) or, equivalently, the aggregate supply function converges to the demand function. As $\sigma^2_e/\sigma^2_\theta \to \infty$ the SFE also collapses and there is a discontinuity in the limit when $\rho \neq 0$.\(^{25}\)

There are particular parameter combinations when $\rho > 0$ for which the scarcity and informational effects balance and sellers set a zero weight ($c = 0$) on public information. In this case sellers do not condition on the price and the model reduces to the Cournot model where sellers compete in quantities. However, in this particular case, when supply functions are allowed, not reacting to the price (public information) is optimal. Figure 1 depicts the change in the equilibrium supply function as $\rho$ goes from $-(n-1)^{-1}$ to 1 for $s_i = \overline{\theta} : X(\overline{\theta}, p)$.

\(^{24}\) From the expression for the weight $a$ of private information in the strategy of a seller and the fact that $d$ decreases in $c$ we have that $a$ decreases in $\rho$ and in $\sigma^2_e/\sigma^2_\theta$ when $\rho \geq 0$.

\(^{25}\) When $\sigma^2_e/\sigma^2_\theta \to \infty$ we have that $a \to 0$ and $c \to \hat{c}$, with $\hat{c} = c'$ for $\rho = 0$ and $c' > \hat{c}$ ($c' < \hat{c}$) for $\rho > 0$ ($\rho < 0$). (See Claim A.2 in the Appendix.) However, the equilibrium in the limit economy with $\sigma^2_e/\sigma^2_\theta \to \infty$ (even when $\rho = 1$) is given by $X(\rho) = c'(p - \overline{\theta})$ since $E[\theta | s, \rho] = \overline{\theta}$, and therefore it coincides with the limit of the SFE as $\sigma^2_e/\sigma^2_\theta \to \infty$ only when $\rho = 0$. There is a discontinuity in the equilibrium correspondence when $\rho \neq 0$. This discontinuity disappears when there is noise in the demand function (see in Section 5.2).
Figure 1. The SFE $X(\bar{\theta}, p)$ as $\rho$ goes from $-(n-1)^{-1}$ to 1.

Constant marginal costs\textsuperscript{26} ($\lambda = 0$): If $n-2 \geq M$ there is no SFE (Proposition 1 (ii)) and the slope of supply degenerates to competitive ($c \to \infty$) as $\lambda \to 0$. However, whenever adverse selection is important enough $n-2 < M$ then as $\lambda \to 0$, $c \to c_0$ and there is a SFE with slope $c_0$ (negative when $n < M$). (See Claim A.3 in the Appendix.) In short, for high enough adverse selection sellers have market power even with constant returns.

Our results are related to the winner’s curse in common value auctions (Milgrom and Weber (1982)): A bidder refrains from bidding aggressively because winning conveys the news that the signal the bidder has received was too optimistic (the highest signal in the pool). Bidders shade their bid more, to protect against the winner’s curse, the less precise their signals are (see Reece (1978)). In our model a seller refrains from competing aggressively with its supply function because a high price conveys the bad news that costs are high and the more so the less precise his signal is. However, in the typical auction model sellers bid for a unit of a good while in our model sellers compete in schedules to

\textsuperscript{26} This case approximates classical multi-unit auction environments.
fulfill a demand for a divisible good and therefore the analogy works with respect to adverse selection but not necessarily with respect to market power. 27

The results are also reminiscent of asymmetric information models where traders submit steeper schedules to protect themselves against adverse selection. 28 Biais et al. (2000) in a common value environment in a discriminatory auction show that adverse selection reduces the aggressiveness of competition in supply schedules of risk neutral uninformed market makers, facing a risk averse informed trader who is subject also to an endowment shock. At the unique equilibrium in convex supply schedules the outcome is of imperfect competition because marginal prices are increasing with the size of trade as market makers protect themselves against informed trading. The latter combined with the optimal response of the informed agent determines a residual demand curve with finite elasticity for every market maker. This imperfect competition result disappears in a pure private value environment where there is no asymmetric information about the value of the asset and adverse selection arises only out of the idiosyncratic endowment shock to the trader. Then marginal prices need no longer be increasing in the amount traded to reflect the informational content of trade. In both Biais et al. (2000) and in our paper private information generates market power. 29 In our model in the pure private value case ($\rho = 0$) there is some market power provided that $\lambda > 0$ and it vanishes when $\lambda \to 0$.

2.4 Competitiveness

The competitiveness of a market is usually measured in terms of absolute and relative margins over marginal costs which are closely related to the perceived elasticity of the

---

27 Our results are perhaps more closely related to the generalized winner’s curse or “champion’s plague” pointed out in Ausubel (2004) for multi-unit auctions according to which, and translated in our context, the expected cost of a bidder conditional on being allocated a larger quantity is larger than with a smaller quantity.

28 I will discuss the precise relationship with the Kyle (1989) model when introducing noise and demand schedule competition (Section 5.4).

29 However, their framework is very different from ours: In their paper the competing market makers/sellers are uninformed while privately informed in ours; the monopsonistic informed buyer selects quantities in the posted schedules while we have a passive competitive demand; the buyer has liquidity shock while there is no noise in our demand; and the auction is discriminatory while ours is uniform. As we will see in Section S.1 of the Supplement (Vives (2011a)) the behavior of a large market is quite different in both models.
residual demand of a seller. For seller \( i \) the residual demand is
\[
\beta^{-1} (\alpha - p) - \sum_{j \neq i} X(s_j, p),
\]
with elasticity \( \hat{\eta}_i = p/dx_i \). The (absolute value of the) slope of residual demand is
\[
d^{-1} = \beta^{-1} + (n-1)c.
\]
From the equilibrium F.O.C. we have that
\[
p - E_i[MC_i] = dx_i,
\]
where \( E_i[MC_i] \equiv E[\theta_i|s] + \lambda x_i \) is the (interim) expected marginal cost of seller \( i \). In Lerner index form,
\[
\frac{p - E_i[MC_i]}{p} = \frac{1}{\hat{\eta}_i}.
\]
A similar relation holds for the margin over average (interim) expected marginal cost \( E_n[MC_n] = n^{-1} \sum_{i=1}^n E_i[MC_i] \),
\[
p - E_n[MC_n] = d\bar{x}
\]
and, correspondingly, for the aggregate (interim) Lerner index,
\[
\frac{p - E_n[MC_n]}{p} = \frac{1}{(\beta^{-1} + (n-1)c)\beta n\eta} = \frac{d}{\beta n\eta},
\]
where \( \eta = p/(\beta n\bar{x}) \) is the elasticity of demand. It follows that
\[
p = E[\bar{\theta}|\bar{s}] + (d + \lambda)\bar{x}.
\]

Three important benchmarks for rivalry are perfect competition, Cournot and collusion. If sellers are price takers and act with full information \( s = (s_1, \ldots, s_n) \) then \( p - E_i[MC_i] = 0 \) and \( p - E_n[MC_n] = 0 \), and this corresponds to the case when \( c = \infty \) and \( d = 0 \).

The case \( c = 0 \) corresponds to a Bayesian Cournot equilibrium, where seller \( i \) sets a quantity contingent only on his information \( \{s_i\} \), and the aggregate (interim) Lerner index is \( (n\eta)^{-1} \). The supply function and the Cournot equilibrium (and allocations) coincide when \( M = n(1 + \lambda \beta^{-1})^{-1} \), in which case \( c = 0 \) and \( d = \beta \). When \( c > 0 \), we are in the

\[30\] Noting that \( n^{-1} \sum_{i=1}^n E[\theta_i|s] = E[\bar{\theta}|s] = E[\bar{\theta}|ar{s}] \), the latter equality holding since \( \bar{s} \) is a sufficient statistic for \( s \) in relation to \( \bar{\theta} \).

\[31\] There is a unique Bayesian Cournot equilibrium and it is linear (see Proposition S.1 in the Supplement (Vives (2011a))).
usual case in which the supply function equilibrium has positive slope and is between the Cournot and the competitive outcomes (e.g. Klemperer and Meyer (1989) when uncertainty has unbounded support and with full information, in which case \( c = c^f > 0 \)). However, when \( c < 0 \) the aggregate (interim) Lerner index is larger than the Cournot level.

If sellers were to collude with full information (share the signals \( s = (s_1, ..., s_n) \) and maximize joint profits) it is easy to see that we would obtain the usual collusive (monopoly) Lerner formula,

\[
\frac{p - E_s[MC_s] - \eta}{p} = \frac{1}{\eta}.
\]

What is surprising, as we will show, is that as \( \rho \) ranges from \(- (n-1)^{-1} \) to 1 we have that \( d \) ranges from 0 to \( \beta n \), and, correspondingly, the price ranges from competitive to collusive. The following proposition states the competitiveness-related results plus a volatility result.

\textbf{Proposition 2.}\ Let \(- (n-1)^{-1} < \rho < 1 \) and \( 0 < \sigma^2_{\varepsilon}/\sigma^2_{\theta} < \infty \), then at the SFE:

(i) The slope of equilibrium supply is steeper (\( c \) smaller) with increases in \( \rho \) and \( c \) ranges from \( \infty \) to \(- 1/\beta n \), and \( d \) from 0 to \( \beta n \), as \( \rho \) ranges from \(- (n-1)^{-1} \) to 1. When \( \rho > 0 \) (\( \rho < 0 \)), \( c \) decreases (increases) with \( \sigma^2_{\varepsilon}/\sigma^2_{\theta} \).

(ii) As \( \rho \) ranges from \(- (n-1)^{-1} \) to 1 the price ranges from competitive to collusive. When \( c > 0 \) (\( c < 0 \)) the price is smaller (larger) than the Cournot level. When \( \rho > 0 \) (\( \rho < 0 \)) the price is larger (smaller) than the full information level.

(iii) The expected price \( \bar{p} \) and margin \( \bar{p} - E[MC_s] = dE[\bar{x}] \) are increasing in \( \rho \) and \( \sigma^2_{\varepsilon}/\sigma^2_{\theta} \) (when \( \rho > 0 \)), and with \( n^{-1} \) (for \( c > 0 \) when \( \rho \geq 0 \)). They are decreasing in \( \sigma^2_{\varepsilon}/\sigma^2_{\theta} \) when \( \rho < 0 \).

(iv) Price volatility \( \text{var}[p] \), when \( \rho > 0 \), decreases with \( \sigma^2_{\varepsilon} \) and increases with \( \sigma^2_{\theta} \).
It is remarkable that sellers may approach aggregate collusive margins in a one-shot noncooperative equilibrium because of informationally-induced market power. Let us recall that at the full information equilibrium (corresponding to $\rho = 0$), indicating the pure market power distortion, the aggregate Lerner index would equal $d' / (\beta \eta)$. As $\rho \to 1$ the private information distortion becomes more severe and sellers protect themselves by increasing the slope of their supplies and become less and less aggressive. Furthermore, as $\rho \to -(n-1)^{-1}$ the increased “favorable” selection implies that market power is reduced and eliminated in the limit.

The explanation of the result is as follows. The aggregate margin and output tend to the collusive level, maximal market power, because as $\rho \to 1$, $d \to \beta n$, the aggregate supply function converges to the demand function and the market collapses. This is precisely the case in which the slope of residual demand for an individual seller is collusive since sellers tend to produce the same (as we will see in Section 3.1, despite that there is some productive inefficiency as long as $\rho < 1$) and as $\rho \to 1$, $\frac{1}{\eta_0} \to \frac{1}{\eta}$ and

$$\frac{p - E_e[MC_s]}{p} \to \frac{1}{\eta}.$$  

This would not happen if equilibrium were to exist for $\rho = 1$. Indeed, with noisy demand (Proposition 8) equilibrium exists even if $\rho = 1$ and the aggregate margin is never fully collusive. (In a similar vein, as $\rho \to -(n-1)^{-1}$ the competitive outcome obtains as the equilibrium collapses also since $c \to \infty$, see Figure 1, but with demand uncertainty it does not and the limit then is not fully competitive.)

**Remark 1**: It is worth noting that the distortion $d$ may increase with $n$ when $c < 0$.\(^{32}\)

This does not happen with full information -then $c^f (d^f)$ is increasing (decreasing) in $n$

---

\(^{32}\) For example, $d(n = 3) > d(n = 2)$ with parameters $\beta = \lambda = 1$, $\sigma^p = \sigma^s = 1$ when $\rho$ is close to 1. See Figure S.1a in Section S.3 of the Supplement (Vives (2011a)) which contains more results and details of the simulations.
(see Claim A.4 in the Appendix)- and it will not happen either when demand is replicated with the number of sellers (see Section 4).

3. Welfare analysis

In order to assess the welfare loss at the SFE we provide an outcome-based characterization of the equilibrium and characterize the deadweight losses. At the full information equilibrium sellers have market power and there is no private information. There is a welfare loss due to market power. At the SFE there is an additional welfare loss due to private-information-induced market power. I show also that the efficient outcome can be implemented with a price-taking supply function equilibrium and how subsidies can implement the efficient allocation.

3.1 A characterization of the SFE outcome and welfare

Let, \( t_i = E[\theta_i | s] \), \( i = 1, \ldots, n \) and \( t = (t_1, \ldots, t_n) \) be the predicted values with full information \( s \). The strategies at a SFE, where \( E[\theta_i | s, p] = t_i \), induce outcomes as a function of the realized vector of predicted values \( t \): \( (x_i(t))_{i=1}^n \) and \( p(t) \). It is easy to see then than the outcome at the SFE maximizes a distorted surplus function with common information \( t \):

\[
\max_{(x_i)_{i=1}^n} \left\{ E[TS | t] - \frac{d}{2} \sum_{i=1}^n x_i^2 \right\}
\]

where \( d \) is the equilibrium SFE parameter (Proposition 1). That is, the market solves the surplus maximizing program with a distorted cost function which represents both higher total and marginal costs:

\[
\hat{C}(x_i, \theta_i) \equiv C(x_i, \theta_i) + \frac{d}{2} x_i^2 .
\]

The result follows since the (sufficient) F.O.C. of the distorted planning problem are:

\[33\] The SFE allocation would be obtained by price-taking sellers with distorted costs functions \( \hat{C}(x_i, \theta_i) \) and full information \( t \). However in this case, supply functions would always be upward sloping, \( x_i = \left( p - E[\theta_i | t] \right)/(d + \lambda) \), since there is no informative role for the price to play.
\[
p - E[\theta | t] - (d + \lambda)x_i = 0, \ i = 1, \ldots, n, 
\]
which are identical to those of the SFE since: \( E[\theta | s, p] = E[\theta | t] \). Similarly, the full information supply function equilibrium can be obtained as the solution to a distorted planning program replacing \( d \) by \( d' \). It is clear that the full (shared-) information efficient allocation obtains setting \( d = 0 \). The implied allocation is symmetric (since the total surplus optimization problem is strictly concave and sellers and information structure are symmetric).

We can consider an SFE allocation parameterized by \( d \) for a given realization of predicted values \( t \), \( (x_i(t; d))_{i=1}^n \). The deadweight loss \( E[DWL | t] \) at the SFE is the difference between total surplus at \( (x_i(t; d))_{i=1}^n \) and at the efficient allocation \( (x_i(t; d = 0))_{i=1}^n \). The wedge \( d > 0 \) induces both distributive/productive and aggregate/allocation inefficiency. Distributive inefficiency refers to an inefficient distribution of sales/production of a given aggregate (average) quantity \( \bar{x} \). Sellers minimize distorted costs \( \hat{C}(x_i, \theta_i) \) with \( d > 0 \), equivalent to a fictitious more convex technology, and the choices of individual quantities are biased towards too similar sales: \( x_i - \bar{x} = (\bar{t} - t)(d + \lambda)^{-1} \) while cost minimization would require letting \( d = 0 \). Aggregate inefficiency refers to a distorted level of average quantity while producing in a cost-minimizing way. Note that average quantity \( \bar{x}(t; d) = (\alpha - \bar{t})/(\beta n + \lambda + d) \) is decreasing in \( d \). The impact of the distortion on profits is also of interest. An increase in \( d \) increases margins but also productive inefficiency with an a priori ambiguous impact on profits. The following proposition states the results.

**Proposition 3.** Consider an allocation parameterized by \( d \) for a given realization of predicted values \( t \). Then

(i) both aggregate and distributive inefficiency, and therefore \( E[DWL | t] \), are increasing in \( d \); and
(ii) average profits increase in $d$ for $d$ small and decrease in $d$ for $d$ close to $\beta n$.

The intuition for the result (i) should be clear since increases in $d$, for a given realization of predicted values, reduce average output and bias individual outputs towards excessive similarity. In regard to result (ii) when the distortion is small increasing $d$ increases average profits by increasing margins more than productive inefficiency while the opposite happens when the distortion is large.

Do the welfare results extend when averaging over predicted values, that is, when taking unconditional expectations? From Proposition 3 we have that for given predicted values $t$ both types of inefficiency increase with $d$, and therefore with $\rho$ and $\sigma_\varepsilon^2$, but changes in those parameters change the probability distribution over $t$. The (expected) deadweight loss ($E[DWL]$) at the SFE is the difference between expected total surplus at the efficient allocation ($ETS^*$) and at the SFE ($ETS$). Let $x_i = x_i(t; d)$ and $x_i^* = x_i(t; 0)$. It can be checked that

$$E[DWL] = n \left( (\beta n + \lambda) E\left[ (\bar{x} - \bar{x}^o)^2 \right] + \lambda E\left[ (u_i - u_i^o)^2 \right] \right) / 2,$$

with $u_i \equiv x_i - \bar{x}$ and $u_i^o \equiv x_i^o - \bar{x}^o$, where the first term corresponds to aggregate inefficiency and the second to distributive inefficiency. The following proposition takes into account the averaging effect and characterizes deadweight losses.

**Proposition 4.**

(i) Expected aggregate inefficiency increases always in $\rho$ (if $\sigma_\varepsilon^2 > 0$) while it may increase or decrease in $\sigma_\varepsilon^2$. If $\rho \leq 0$ then it decreases in $\sigma_\varepsilon^2$.

(ii) Expected distributive inefficiency may increase or decrease in $\rho$ and in $\sigma_\varepsilon^2$. If $\sigma_\varepsilon^2 = 0$ then it decreases in $\rho$ and if $\rho \leq 0$ then it decreases in $\sigma_\varepsilon^2$. 


Expected profits, when $\sigma^2 > 0$, converge to the collusive level as $\rho \to 1$ and to the competitive level when $\rho \to -(n-1)^{-1}$. If $\sigma^2 = 0$ then they decrease in $\rho$ and if $\rho \leq 0$ then they decrease in $\sigma^2$.

Remark 2: With respect to the welfare loss induced by full information market power: Both expected aggregate and distributive inefficiency decrease in $\sigma^2$, and expected aggregate (distributive) inefficiency increases (decreases) in $\rho$.

The intuition for the results is as follows. (i) The aggregate inefficiency term $E\left[\left(\bar{x} - \bar{x}^o\right)^2\right] = \left((\beta n + \lambda)^{-1} - (\beta n + \lambda + d)^{-1}\right)^2 E\left[\left(\alpha - \bar{t}\right)^2\right]$ increases in $d$ and in the variance of the prediction $\bar{t} = E[\hat{\theta}|\hat{s}]$, $\text{var}[\bar{t}]$. Increases in $\rho$ increase both (if $\sigma^2 > 0$) while increases in $\sigma^2$ decrease $\text{var}[\bar{t}]$ and this effect may dominate if $\rho$ is small. (ii) The distributive inefficiency term $E\left[\left(u_i - u_i^o\right)^2\right] = \left(\lambda^{-1} - (\lambda + d)^{-1}\right)^2 E\left[\left(t_i - \bar{t}\right)^2\right]$ increases in $d$ and in $E\left[\left(t_i - \bar{t}\right)^2\right] = \text{var}[t_i - \bar{t}]$. The non-monotonicity with respect to $\rho$ (if $\sigma^2 > 0$) and $\rho^2$ (if $\rho > 0$) follows since $\text{var}[t_i - \bar{t}]$ decreases in $\rho$ (if $\sigma^2 > 0$) and $\sigma^2$ (if $\rho > 0$). Indeed, with $\rho = 1$ or $\sigma^2 = \infty$ there would be no distributive inefficiency. The results when $\rho \sigma^2 = 0$ follow since then $d$ is independent of $\rho$ and $\sigma^2$ and only the averaging effect is present, and when $\rho < 0$ then $d$ is decreasing in $\sigma^2$.

(iii) When $\rho \to 1$ and $\sigma^2 > 0$ we know (from Proposition 2) that $d \to \beta n$ and $\bar{x}(t,d) \to \bar{x}^o \equiv \bar{x}(t,\beta n)$, the average collusive output. Furthermore, as $\rho \to 1$, firms to produce the same and productive inefficiency at the SFE vanishes, and expected profits converge to the collusive level. The same result holds replacing collusive by competitive when $\rho \to -(n-1)^{-1}$.

---

34 The same applies for the welfare loss induced by standard market power since $d = d'$ is independent of $\rho$ and $\sigma^2$. 

22
3.2 Simulations

In the central scenario of the simulations (with $\rho \in [0,1]$) increases in $\rho$ or $\sigma^2_\epsilon$ increase the deadweight loss at the SFE ($E[DWL] = ETS^\circ - ETS$) - see Figure 2. However, increasing $\rho$ may decrease $E[DWL]$ when $\sigma^2_\epsilon$ is small for a range of $\rho$ bounded away from 1, and increasing $\sigma^2_\epsilon$ may decrease $E[DWL]$ when $\rho$ is small. Furthermore, expected profits $E[\pi_t]$ increase in $\rho$ or $\sigma^2_\epsilon$ provided $\rho$ or $\sigma^2_\epsilon$ are not too close to 0. Otherwise $E[\pi_t]$ may decrease in $\rho$ or $\sigma^2_\epsilon$, and this will tend to be so for $\sigma^2_\epsilon$ large.

The outcome of the simulations performed suggests thus that the results of Proposition 3(i) derived for given predicted values of cost parameters extend to averaging over those values provided that $\rho$ and $\sigma^2_\epsilon$ are not too small. This is so since for given $t$ increases in $\rho$ or $\sigma^2_\epsilon$ increase both aggregate and distributive inefficiency. When averaging over predicted values distributive inefficiency decreases with increases in $\rho$ or $\sigma^2_\epsilon$ and the effect overwhelms the impact given $t$ when $\sigma^2_\epsilon$ (or $\rho$) are small enough. With respect to expected profits the results of Proposition 3(ii) do not extend to the ex ante situation since now increasing $\rho$ or $\sigma^2_\epsilon$ increases the expected margin and although it increases also the distortion $d$, it reduces the expected distributive inefficiency because the predictions of sellers are more aligned. However, when either $\rho$ or $\sigma^2_\epsilon$ is small distributive inefficiency may weigh more.

**Remark 3:** The deadweight loss due to private-information-induced market power $ETS^f - ETS$ (which equals $E[DWL]$ minus the deadweight loss in the full information equilibrium $ETS^\circ - ETS^f$) increases in $\rho$ or $\sigma^2_\epsilon$ ($E[DWL]$) tends to increase with increases in $\rho$ or $\sigma^2_\epsilon$ while $ETS^\circ - ETS^f$ diminishes with $\sigma^2_\epsilon$ always and distributive

---

35 See Section 3.1 in the Supplement (Vives (2011a)) for details and more results of the simulations.

36 Recall that $E[\pi_t]$ decrease in $\rho$ when $\sigma^2_\epsilon = 0$ and in $\sigma^2_\epsilon$ when $\rho = 0$ (Proposition 4(iv)). See Figures S.2 and S.3 in the Supplement (Vives (2011a)) for illustrations of the simulations.
inefficiency -which tends to dominate- also diminishes with $\rho$). $ETS' - ETS$ may decrease in $\sigma^2_\varepsilon$ when $\rho$ is small.

Figure 2. $E[DWL] = ETS^\sigma - ETS$ as a function of $\rho$ and $\sigma^2_\varepsilon$ (with parameters $\beta = \lambda = 1$, $\sigma^2_\theta = 1$, $n = 4$, $\alpha = 200$, $\overline{\theta} = 20$).

Supply function versus Cournot. For $\sigma^2_\varepsilon$ or $\rho$ small, sellers at the supply function market act with full information, $c > 0$ and have less market power than in the (Bayesian) Cournot equilibrium, where sellers do not act with full information. For larger $\rho$ and $\sigma^2_\varepsilon > 0$, $c < 0$ and sellers in the supply function market have more market power and this may dominate the information effect. Simulations suggest that for parameters for which $c > 0$ at the SFE, the supply function market attains a higher expected total surplus than the Cournot market and, for $n$ not too large, the opposite happens when $c < 0$ (e.g. for $\rho$ close to 1).37 (See Sections S.2, S.3.3 and Figure S.5 in the Supplement (Vives (2011a)) When a supply function market is modeled, for convenience, à la Cournot a bias is introduced, overestimating the welfare loss with respect to the actual supply function mechanism on two counts when supply functions slope upwards: excessive market power and lack of information aggregation. When the equilibrium supply function

---

37 When parameters are such that $c = 0$ the SFE and (Bayesian) Cournot allocations coincide and therefore they are both equally efficient.
slopes downwards the Cournot market underestimates market power and then the Cournot market may in principle under- or over-estimate the deadweight loss in relation to supply function competition.

3.3 Price-taking equilibrium, efficiency, and optimal subsidies
If sellers would act as price-takers and had full information the allocation would be (full information) efficient. It is easy to see that the efficient full information allocation can be implemented by a symmetric price-taking linear Bayesian supply function equilibrium (price-taking SFE for short, denoted with a superscript \( PT \)). This is an equilibrium where sellers do not perceive the influence that their supply decisions have on prices but still condition on their private signals and try to learn from prices. The F.O.C. for a price-taking SFE are the same as in the proof of Proposition 1 letting \( d = 0 \):

\[
p = E[\theta | s_i, p] + \lambda x_i \text{ for } i = 1, \ldots, n,
\]

yielding a supply function \( X^{PT}(s, p) = \frac{p - E[\theta | s_i, p]}{\lambda} \). As before, in equilibrium \( p \) reveals \( s_\ast, E[\theta | s_i, p] = E[\theta | s_i, \tilde{s}] \), and the price-taking equilibrium implements the efficient solution since sellers have full information and act competitively.

From the previous analysis it may be conjectured that first best efficiency may be restored by a quadratic subsidy \( \kappa x_i^2 / 2 \) that “compensates” for the distortion \( dx_i^2 / 2 \) and induces sellers to act competitively.\(^{38}\) The question is whether we can find a \( \kappa \) (with \( \kappa > 0 \) for a subsidy) such that \( \lambda - \kappa + d(\kappa) = \lambda \), where \( d(\kappa) \equiv (\beta^{-1} + (n-1)c(\lambda - \kappa))^{-1} \) is the (endogenous) distortion when the slope of marginal cost is \( \lambda - \kappa \). That is, whether we can find a solution to the fixed-point equation \( d(\kappa) = \kappa \). If we can find such a \( \kappa \), a seller would act effectively as if he was competitive and facing a marginal cost with slope \( \lambda \) and the F.O.C. would be

\(^{38}\) Angeletos and Pavan (2009) provide a thorough analysis of tax-subsidy schemes in quadratic continuum economies with private information and with agents using non-contingent strategies (e.g. of the Cournot type). In their model, however, there is no learning from endogenous public signals and taxes are made contingent on aggregate realizations.
\[ p - E[\theta | s] - (d + \kappa - \kappa) x_i = p - E[\theta | s] - \lambda x_i = 0. \]

The following proposition states the results.

**Proposition 5.** Let \(- (n-1)^{-1} < \rho < 1 \) and \( \sigma_{\epsilon}^2 / \sigma_\rho^2 < \infty \). Then:

(i) There is a unique price-taking SFE and the equilibrium implements the efficient allocation. The slope of supply is given by \( c^{PT} = (\lambda^{-1} - (\beta n)^{-1} M)/(M + 1) \), which is decreasing with \( M \) and \( \lambda \).

(ii) There is an optimal quadratic subsidy \( \kappa^* x_i^2 / 2 \), \( \kappa^* = (\beta^{-1} + (n - 1) c^{PT} (\lambda))^{-1} \), which implements price-taking behavior. Implementation need not be unique if adverse selection is severe: \( n - 2 < M \). The optimal subsidy \( \kappa^* \) increases with \( \lambda \) and \( \rho \), and it increases (decreases) with \( \sigma_{\epsilon}^2 \) when \( \rho > 0 \) (\( \rho < 0 \)).

**Remark 4:** In contrast to result (i) at a price-taking Bayesian quantity-setting equilibrium there is typically a welfare loss because of lack of information aggregation.\(^{39}\)

**Remark 5:** We have that \( c^{PT} > 0 \) for \( M \) small (and negative a fortiori) and \( c^{PT} < 0 \) for \( M \) large. The price-taking supply function will coincide with the marginal cost schedule only when there is no learning from prices (that is, when \( M = 0 \), in which case both schedules boil down to \( p = E[\theta | s] + \lambda x_i \)).\(^{40}\) The supply function of a seller in the price-taking equilibrium is always flatter than the supply function in the strategic equilibrium since \( d > 0 \) and \( c^{PT} - c = (\lambda^{-1} - \frac{d + \lambda}{M + 1}) > 0. \(^{41}\)

---

\(^{39}\) See Section S.2 in the Supplement (Vives (2011a)).

\(^{40}\) As in the strategic case the supply function of a seller converges to the per capita seller demand function as \( \rho \to 1 \), and \( a^{PT} \to \infty \) as \( \rho \to -(n-1)^{1/2} \).

\(^{41}\) Sellers are more cautious responding to their private signals when they are strategic since they take into account the price impact coming from the amount sold as well as the potential informational leakage from their actions: \( a^{PT} - a > 0 \). By the same token, given that \( d > d' \) we have that \( c^{PT} > c' > c \) and \( a^{PT} > a' > a \).
To understand result (ii) note that we have to find a $\kappa$ such that $c(\lambda - \kappa) = c^{PT}(\lambda)$, and this will yield $d(\kappa) \equiv (\beta^{-1} + (n-1)c^{PT}(\lambda))^{-1}$. When $n - 2 \geq M$, and the adverse selection problem is moderate, we are in the situation depicted in Figure 4a and we can find a subsidy $\lambda - \kappa > 0$ to implement price-taking behavior. However, when $n - 2 < M$ we are in the situation depicted in Figure 4b and we may need to induce effective increasing returns: $\lambda - \kappa < 0$. However, for negative slopes of marginal costs there are two linear equilibria (with slopes of supply $c_2$ and $c_1$ -see Lemma A.1 in the Appendix). This means that we can find an optimal quadratic subsidy always but the implementation of the efficient allocation is unique only when there is no need to induce negative slopes of effective marginal costs. When $n - 2 < M$ and $c^{PT}(\lambda) > c_0$ there is one equilibrium that implements price-taking behavior with $\lambda - \kappa^* < 0$. For example, for $\hat{\lambda}$ such that $c^{PT} > \hat{c}$ ($\hat{\lambda}''$ in Figure 4b) we need to choose the higher $c_1$ equilibrium, while for lower $\lambda$ with $c_0 < c^{PT} < \hat{c}$, ($\hat{\lambda}'$ in Figure 4b) we need to choose the lower $c_2$ equilibrium. It is worth noting that, indeed, when $\tau \to -(n-1)^{-1}$ competitive behavior is already approached in the market with no subsidy and, therefore, $\kappa^* \to 0$.

![Figure 4a](image-url)

*Figure 4a*: The equilibrium parameters $c^{PT}$ and $c$ as a function of $\lambda$ and the optimal subsidy $\kappa^*$ when $n - M - 2 \geq 0$ and $M > 0$. 
Figure 4b: The equilibrium parameters $c^{PT}$ and $c$ as a function of $\lambda$ and the optimal subsidy $\kappa$ when $n - M - 2 < 0$ (the case depicted is with $n - M > 0$). When $\lambda = \lambda'$ the optimal subsidy is $\kappa'$ with $c_2(\lambda - \kappa') = c^{PT}(\lambda)$ and when $\lambda = \lambda''$ the optimal subsidy is $\kappa''$ with $c_1(\lambda - \kappa'') = c^{PT}(\lambda)$.

4. Convergence to price-taking behavior in large markets

In order to study whether (and if so how fast) the inefficiency of supply function equilibria disappears in large markets we consider replica markets where the numbers of sellers and buyers grow at the same rate $n$. More precisely, suppose that there are $n$ buyers, each with quasilinear preferences and benefit function $u(x) = \alpha x - \beta x^2/2$ where $x$ is the consumption level. This gives rise to the inverse demand $P_n(y) = \alpha - \beta y/n$ where $y$ is total consumption. There are $n$ sellers as before. Total surplus is therefore given by $TS = nu(y/n) - \sum_i C(x_i, \theta_i)$ and per capita surplus by $TS/n$. We restrict attention in this section to the case of nonnegative correlation $\rho \in [0, 1]$.

We denote with subscript $n$ the magnitudes in the $n$-replica market. The results we have obtained so far, except possibly comparative statics with respect to $n$, hold replacing $\beta$ by $\beta/n$. The following proposition characterizes the convergence of the SFE to a price-taking equilibrium as the market grows. (See Section S.1 in the Supplement (Vives (2011a)) for the definitions of orders of sequences, further results, proofs and comments.)
As we have seen before the price-taking equilibrium is first best efficient since it aggregates information. We confirm that the efficient outcome is approached as the market becomes large.

Proposition 6. Let \( \rho \in [0,1) \). In the replica market:

(i) As the market grows large the market price \( p_n \) at the SFE converges in mean square to the price-taking Bayesian price \( p_n^{PT} \) at the rate of \( 1/n \).

(ii) The deadweight loss at the SFE \( (ETS_n^o - ETS_n) \) is of the order of \( 1/n^2 \).

A large market approaches efficiency in prices at a rate \( 1/n \), which is the same as the usual rate under complete information. This is a statement that bid shading is of the order of \( 1/n \). It follows from the fact that the distortion \( d_n = (n\beta^{-1} + (n-1)c_n)^{-1} \) is of order \( 1/n \), and both equilibria aggregate information since \( p_n \) and \( p_n^{PT} \) reveal the average signal \( \bar{s}_n \).

Simulations suggest that both \( d_n \) and \( E[DWL_n]/n \) are monotonically decreasing in \( n \).

Cripps and Swinkels (2006) obtain a parallel result in a double auction environment. The authors consider a generalized private value setting where bidders can be asymmetric and can demand or supply multiple units. Under some regularity conditions (and a weak requirement of “a little independence” where each player’s valuation has a small idiosyncratic component), they find that as the number of players grows (say that there are \( n \) buyers and \( n \) sellers) all nontrivial equilibria of the double auction converge to the competitive outcome and inefficiency vanishes at the rate of \( 1/n^{2-\chi} \) for any \( \chi > 0 \).

As in Kyle (1989) we could ask what happens when the total amount of precision available to agents is fixed as the market grows large. Then as \( n \) tends to infinity \( \tau \rightarrow 0 \). In this case it is easy to see that \( M \rightarrow \infty \) as \( n \rightarrow \infty \) and, as when \( \rho \rightarrow 1 \), the

\[ \kappa_n^* = \left( \beta^{-1}n + (n-1)c_n^{PT}(\lambda) \right)^{-1} \]

can be checked also to be decreasing in \( n \).

42 See Section S.3.2 in the Supplement (Vives (2011a)) for further results and details of the simulations.
supply function of a seller converges to the per capita demand, \( d \to \beta n \), and the equilibrium collapses. In the limit we approach the collusive price (as in Proposition 2) and therefore an extreme form of the non-competitive limit of Kyle.

**Remark 6:** For a given \( \rho > 0 \) and for large enough \( n \) we have always that the supply function equilibrium attains higher surplus than the (Bayesian) Cournot equilibrium. This is so since as \( n \) grows the SFE, but not the Cournot equilibrium, converges to the full information first best. In a large enough market the Cournot model always overestimates the welfare loss since a deadweight loss remains due to private information when \( \rho > 0 \) (while convergence to price taking behavior obtains as in the supply function market).\(^{43}\)

## 5. Extensions

In this section we test the robustness of the results in the context of inelastic demand, demand uncertainty, public signals, and demand schedule competition.

### 5.1 Inelastic demand

The case where an auctioneer demands \( q \) units of the good is easily accommodated letting \( \beta \to \infty \) and \( \alpha / \beta \to q \). Then from the inverse demand we obtain

\[
y = (\alpha - p) / \beta \to q.
\]

Let \( \rho > -(n-1)^{-1} \) and \( \sigma_e^2 / \sigma_\theta^2 < \infty \). It can be checked that there is a unique SFE if only if \( n - 2 - M > 0 \).\(^{44}\) In equilibrium we have that

\[
c = \frac{n - 2 - M}{\lambda (n - 1)(1 + M)} > 0,
\]

and \( d = ((n-1)c)^{-1} \) is decreasing with \( n \). A necessary condition for existence of the SFE

---

\(^{43}\) Proposition 7 holds for the (Bayesian) Cournot equilibrium replacing \( ETS^* \) with the expected total surplus at the price-taking (Bayesian) Cournot equilibrium. See Proposition S.2, and Figure S.7 in the Supplement (Vives (2011a)).

\(^{44}\) The analysis is analogous to the proof of Proposition 1. Now the S.O.C. holds if \( c > 0 \). If \( q = 0 \) we are in a double auction case and there is also a no-trade equilibrium. See Vives (2010) for a presentation of the model with demand bidders facing an inelastic supply.
if $\rho \sigma^2_x \geq 0$ is that $n \geq 3$. If $\rho \sigma^2_x < 0$ then $M < 0$ and there is an equilibrium for $n \geq 2$.

As $M$ increases, $c$ decreases and as $n - 2 - M \to 0$, $c \to 0$ and the SFE collapses. This is because of the combination of adverse selection and market power when demand is inelastic: The supply schedules become too inelastic to sustain a linear equilibrium as $n - 2 - M \to 0$. The market breaks down when traders submit vertical schedules since with vertical residual demand curves traders sometimes would like to force unbounded prices. This is similar to the double auction context of Kyle (1989) in which a linear equilibrium exists only if the number of informed traders is larger or equal than 3 (when there are no uninformed traders). In our more general model with strategic agents facing an elastic demand function from passive buyers the market does not break down since there is always price elasticity from demand. With inelastic demand supply functions are always upward sloping ($c > 0$). This is as in the auction models of Kyle (1989) or Wang and Zender (2002) where demand schedules always have the “right” slope.

The results obtained specialize to the inelastic demand case. We highlight here the differences. With regard to Proposition 2 we have that in the range of existence $n - 2 - M > 0$: (ii) the price is always between the competitive and the Cournot price (since $c > 0$); (iii) expected bid shading increases with $q/n$ ($p - E_n[MC_n] = dq/n$ and $d$ is decreasing with $n$); and (iv) $\var[p]$ increases with $\rho$ and $\sigma^2_\theta$, and decreases with $\sigma^2_x$ (since $\bar{x} = q/n$ and therefore $\var[p] = \var[E[\bar{\theta}^\prime \bar{s}]]$). With regard to propositions 3 and 4, the results for distributive inefficiency apply (there is no inefficiency in the aggregate quantity). With regard to Proposition 5: (i) we have that $c^{PT} = (\lambda (M + 1))^{-1}$ and $c \to \infty$ as $\lambda \to 0$, and (ii) for any $\lambda > 0$ there is always a $\kappa \in (0, \lambda)$ such that $c(\lambda - \kappa) = c^{PT}(\lambda)$ with effective marginal costs with positive slope. This yields $\kappa^* = ((n-1)c^{PT}(\lambda))^{-1} = \lambda (1 + M)/(n-1)$. The subsidy $\kappa^*$ increases with $M$ and $\lambda$, and decreases with $n$.  


5.2 Demand uncertainty
Demand uncertainty can be incorporated easily in the model as long as it follows a Gaussian distribution and enters in an additive way: \( P(y) = \alpha + u - \beta y \) with \( u \sim N(0, \sigma_u^2) \) independent of the other random variables. The analysis of the equilibrium proceeds as in Section 2.1 with now the intercept of residual demand \( I_i \) being informationally equivalent to \( h_i = u + \beta a \sum_{j \neq i} s_j \). The following proposition characterizes the equilibrium (see Section S.5 in the Supplement (Vives (2011a)) for a development, complete statement of results and proofs).\(^{45}\)

**Proposition 8.** Let \( \lambda > 0 \). For any, \( \rho \in \left[-(n-1)^{-1}, 1\right] \), \( \sigma_u^2 > 0 \), and \( \sigma_e^2 \geq 0 \), there exists a SFE. It is given by \( X(s, \rho) = \left(p - E[\theta | s_i, \rho] \right)/(d + \lambda) \) where \( 0 < d < \beta n \). As \( \sigma_u^2 \to 0 \), \( d \) tends to the value of \( d \) in Proposition 1 (where \( \sigma_u^2 = 0 \)) and as \( \sigma_u^2 \to \infty \), \( d \to d' \). The equilibrium is unique if \( \rho \geq 0 \), or \( \rho < 0 \) and \( n > 3 \).\(^{46}\) Then if \( \sigma_e^2 > 0 \), \( d \) increases in \( \rho \); and if \( \rho \sigma_e^2 > 0 \) (resp. \( \rho \sigma_e^2 < 0 \)) \( d \) is decreasing (resp. increasing) with \( \sigma_u^2 \).

**Remark 7:** When \( \lambda = 0 \) equilibrium exists if and only if \( n - 2 - M < 0 \) (that is, we need \( \rho > 0 \) and large enough).

The properties of the equilibrium follow.

The equilibrium is noisy. It is immediate that the price is now informationally equivalent to \( u + \beta n \tilde{s} \), where as before \( a \equiv -\partial X/\partial s_i \), and sellers learn only imperfectly from the price about the average signal \( \bar{s} \). Now we have a noisy linear equilibrium instead of a privately revealing one since sellers have on top of their private signal a noisy estimate of

---

\(^{45}\) Vives (2011c) considers the common value case with demand uncertainty in a limit large market and performs a welfare analysis.

\(^{46}\) Simulations suggest that the equilibrium is unique also when \( \rho < 0 \) for \( n = 2, 3 \).
the average signal. The equilibrium exists even when \( \rho = 1 \) or \( \rho = -(n-1)^{-1} \) since with uncertain demand even for extreme values of \( \rho \) there is aggregate uncertainty. The consequence is that the collusive and competitive cases are not attained when \( \sigma_u^2 > 0 \) when \( \rho = 1 \) or \( \rho = -(n-1)^{-1} \) (respectively).

**Conditions for the noise independence property** (equilibrium independent of \( \sigma_u^2 \)).

(i) When \( \rho \sigma_e^2 = 0 \) (\( M = 0 \)), then the equilibrium does not depend on \( \sigma_u^2 \) and \( d = d' \) (as in Proposition 1 when \( \rho \sigma_e^2 = 0 \)). When \( \sigma_e^2 \rightarrow \infty \) and \( \sigma_u^2 > 0 \) then again the equilibrium is independent of \( \sigma_u^2 \), and \( d \rightarrow d' \) yielding \( X(p) = c'(p - \overline{\theta}) \). This limit is also the equilibrium when \( \sigma_e^2 = \infty \). This is in contrast with the case \( \sigma_e^2 = 0 \) where, as we have seen in footnote 25 (Claim A.2 in the Appendix), there is a discontinuity in the limit when \( \sigma_e^2 \rightarrow \infty \). In these cases there is no relevant asymmetric information, no learning form prices, and the equilibrium is independent of the distribution of demand uncertainty (as in Klemperer and Meyer (1989)).

**Comparative statics.** We have that \( d \) decreases (and \( c \) increases) in \( \sigma_u^2 \) when \( \rho \sigma_e^2 > 0 \) because as there is more noise in the demand the information role of the price and the adverse selection problem are diminished. The supply function is flatter with higher \( \sigma_e^2 \) (a high price need not be such bad news about costs since it may come from a high demand realization). This means that as \( \sigma_u^2 \) increases average output increases and the expected margin over marginal cost decreases. The opposite happens when \( \rho \sigma_e^2 < 0 \) since then there is favorable selection and an enhanced information role of the price is pro-

47 When \( \rho = 1 \) and \( \sigma_e^2 = 0 \) we are in fact exactly in the Klemperer and Meyer (1989) case with linear and uncertain demand, symmetric quadratic costs, and no private information (and we obtain the same equilibrium with slope of supply \( c = c' \)).

48 Bernhardt and Taub (2010) obtain the opposite result in a model with private information about the demand for a homogeneous product in which price is observed with noise. Then as the variance of the noise increases, the information role of the price signal is diminished, a high price signal is less likely to mean a high price, and agents rely less on the price.
competitive. As $\sigma_u^2 \to \infty$ we obtain that $d \to d^f$ since then sellers do not learn from the price and rely only on their private signals. As $\sigma_u^2 \to 0$, $d$ converges to the value in the privately revealing equilibrium of Proposition 1.

As $\rho$ increases, and cost parameters become more correlated, $d$ increases since the weight on the information component of the price $h_i$ in $E[\theta_i|s_i, h_i]$ increases with $\rho$ (see Claim S.1 in the Supplement (Vives (2011a))). We have that $d > d^f$ when $\rho \sigma_\varepsilon^2 > 0$ and $d < d^f$ when $\rho \sigma_\varepsilon^2 < 0$. As $\rho \to 1$, $d$ attains its largest value $\dd < \beta n$ for given $\sigma_\varepsilon^2$, $\sigma_\theta^2$ and $\sigma_u^2$, and the price falls short of the collusive level. As $\rho \to - (n-1)^{-1}$, $d$ attains its smallest value $d > 0$. As $\sigma_u^2 \to 0$, $\dd \to \beta n$ and $d \to 0$.

It is easy to see that $d$ is nonmonotone in $\sigma_\varepsilon^2$ when $\rho > 0$. For $\sigma_\varepsilon^2$ small, increases in $\sigma_\varepsilon^2$ increase the value of the price signal $u + \beta n a \tilde{s}$ (the weight on the information component of the price $h_i$ in $E[\theta_i|s_i, h_i]$) while for $\sigma_\varepsilon^2$ large they diminish it (since the price is not very informative as $a \to 0$ when $\sigma_\varepsilon^2 \to \infty$). This does not matter when $\sigma_u^2 = 0$, since then the price is not noisy and recovers $\tilde{s}$ as long as $a > 0$. The result is that $d$ increases with $\sigma_\varepsilon^2$ for $\sigma_\varepsilon^2$ small and decreases with $\sigma_\varepsilon^2$ for $\sigma_\varepsilon^2$ large.

5.3 Public signal
Suppose that sellers receive a public signal on $\tilde{\theta}$, $r = \tilde{\theta} + \delta$ where $\delta \sim N(0, \sigma_\delta^2)$ and $\text{cov}(\tilde{\theta}, \delta) = 0$ (and $\delta$ independent also of the rest of the random variables in the model). Then linear strategies will be of the form $X(s_i, r, p) = b - as_i - er + cp$. A similar analysis

---

49 Note that with demand uncertainty a collusive seller recovers the value of $u$ from the price by using a supply function. The Lerner condition in the collusive case is exactly as in Section 2.4.

50 Indeed, for $\rho > 0$, $d$ attains its minimum value $d = d^f$ both when $\sigma_\varepsilon^2 = 0$ and when $\sigma_\varepsilon^2 \to \infty$. Furthermore, $d$ increases in $\sigma_\varepsilon^2$ when $\sigma_\varepsilon^2 \leq \sigma_\theta^2$ and eventually it decreases as $\sigma_\varepsilon^2$ grows (see (v) Proposition S.5 (iii) in the Supplement (Vives (2011a))).
as in Section 2.1, with the information set of seller $i$ being now $\{s_i, r, p\}$ or, equivalently, $\{s_i, r, h_i\}$ where $h_i \equiv \beta a \sum_{j \neq i} s_j$, leads to the following proposition (see Section S.6 in the Supplement (Vives (2011a)) for a proof).

**Proposition 9.** Let $-(n-1)^{-1} < \rho < 1$, $\sigma^2 / \alpha^2 < \infty$, and $\sigma^2 > 0$, then there is a unique SFE. It is given by the supply function $X(s_i, p) = (p - E[\theta | s_i, r, p]) / (d + \lambda)$ with the parameters $d$ and $a$ characterized as in Proposition 1 with $c$ the largest solution to the quadratic equation $g(c; Q) = 0$ where

$$Q = \frac{\sigma^2 (n^2 \sigma^2 \rho - \sigma^2 (1 - \rho) (1 + (n-1) \rho))}{(1 - \rho) (n \sigma^2 (\sigma^2 (1 + (n-1) \rho) + \sigma^2) + \sigma^2 \sigma^2 (1 + (n-1) \rho))}.$$  

We have that $1 + Q > 0$. As $\sigma^2 \to \infty$, $Q \to M$ and as $\sigma^2 \to 0$, $Q \to -1$, $c \to +\infty$ and $d \to 0$. If $\sigma^2 > 0$ then $Q$, and therefore $d$, increase with $\sigma^2$ and $\rho$, and $\text{sgn}\{\partial Q / \partial \sigma^2\} = \text{sgn}\{Q\}$.

The equilibrium is privately revealing as in our base case, with the price measurable in the public signal $r$ and in $\tilde{s}, p = (1 + \beta nc)^{-1} (\alpha - \beta nb + \beta n (\epsilon r + a \tilde{s}))$, and $\tilde{x} = (\alpha - E[\tilde{\theta} | \tilde{s}, r]) / (\beta n + \lambda + d)$. Now the public signal provides information on aggregate uncertainty $\tilde{\theta}$ beyond the average signal $\bar{s} = \tilde{\theta} + \tilde{\epsilon}$. For $\sigma^2$ low enough we will have $Q < 0$ even if $\rho > 0$ and $d$ will be reduced below the full information level $d^f$ when sellers know $\tilde{s}$ (but not the public signal). The public signal provides additional information to the collective information of sellers. When $Q < 0$ a high price is good news for the costs of a seller (the weight on the price in $E[\tilde{\theta} | s_i, r, p]$ is negative).

---

51 If the public signal were to be just a noisy version of $\tilde{s}$ then as noise vanishes we would recover the full information distortion $d^f$.

52 The coefficient of $h_i$ in $E[\theta_i | s_i, r, h_i]$ equals $(d + \lambda) Q / \beta n$ (see Claim S.3 in the Supplement (Vives (2011a))). To fix the intuition consider the case $\rho = 0$. Then for seller $i$ the public signal $r$ provides a noisy signal about $\theta_i$: $rn = \theta_i + (\sum_{j \neq i} \theta_j + n \delta)$. The price provides a noisy signal about
Note that \( Q < 0 \) when \( \rho \sigma_\delta^2 \) is positive and small or negative. When \( \rho < 0 \) there is favorable selection as in the case where there is no public signal.

The index of adverse selection is now \( Q \) instead of \( M \). It follows that the effects of changes in \( \rho \) on \( c \) and \( d \) are the same as in the case \( \sigma_\delta^2 = \infty \) since \( Q \) is increasing in \( \rho \).\(^{53}\) When \( \rho \to 1 \) the distortion \( d \to \beta n \) (since \( Q \to \infty \)) and we approach collusive pricing (as in the case when there is no public signal). Similarly as in the base case, and for the same reasons, we have that as \( \rho \to 1 \) the linear equilibrium collapses. When \( \rho \to -(n-1)^{-1} \) we approach also the competitive outcome (since then \( Q \to -1 \)).

Increases in \( \sigma_\varepsilon^2 \) will decrease \( c \) if \( \rho \sigma_\delta^2 \) is large, but if \( \rho \sigma_\delta^2 \) is small or negative then \( c \) will increase with \( \sigma_\varepsilon^2 \). This is so since \( \text{sgn}\{\partial Q / \partial \sigma_\varepsilon^2\} = \text{sgn}\{Q\} \) and when \( \rho \sigma_\delta^2 \) is small (but positive) or negative, \( Q < 0 \) and \( d < d^f \). Then increases in \( \sigma_\varepsilon^2 \) reduce \( Q \) and increase \( c \) since the information role of the price is enhanced and a high price is good news.\(^{54}\)

As the precision of the public signal improves (\( \sigma_\delta^2 \) decreases), the adverse selection problem diminishes, and the slope of equilibrium strategies as well as the distortion are lowered. In fact, when \( \sigma_\delta^2 \to 0 \) the distortion is eliminated, \( d \to 0 \) (since \( Q \to -1 \)). Note that in the limit as \( \sigma_\delta^2 \to 0 \), there is no equilibrium (indeed, \( c \to +\infty \)).\(^{55}\)

\[
\sum_{j \neq i} \theta_j \quad \text{and therefore helps reading the public signal} \ r. \ \text{When the seller sees a high price then it infers that} \ \sum_{j \neq i} \theta_j \ \text{is high and therefore} \ \theta_i \ \text{low, the more so the less noisy is the public signal} \ (\sigma_\delta^2 \ \text{low}). \ \text{A similar intuition applies when} \ \rho \sigma_\delta^2 \geq 0 \ \text{is small, and this explains why then} \ Q < 0 \ \text{and a high price is good news. As} \ \rho \ \text{increases eventually a high price becomes bad news.}
\]

\(^{53}\) The same applies in relation to \( \sigma_\delta^2 \) for \( \rho \) not too negative since then \( Q \) is decreasing in \( \sigma_\delta^2 \) (\( \text{sgn}\{\partial Q / \partial \sigma_\delta^2\} = -\text{sgn}\{n^2 \sigma_\delta^2 \rho + \sigma_\varepsilon^2 (1 + (n-1) \rho)\} \), see the proof of Proposition 9).

\(^{54}\) Indeed, when \( \sigma_\delta^2 > 0 \) the absolute value of the weight on \( h \) in \( E[\theta \mid s, r, h] \) increases in \( |\rho| \) and \( \sigma_\varepsilon^2 \) (see Claim S.2 in Section S.6 of the Supplement (Vives (2011a))).

\(^{55}\) When \( \sigma_\delta^2 = 0 \), \( E[\tilde{\theta} \mid \hat{s}, r] = r = \tilde{\theta} \) (since \( \hat{s} = \tilde{\theta} + \tilde{e} \)) and the candidate equilibrium price does not depend on the average signal. This implies that at a symmetric equilibrium sellers do not put any
5.4 Demand schedule competition

The model can be restated in terms of competition among buyers of an asset of unknown ex-post average value \( \tilde{\theta} = \left( \sum_{i=1}^{n} \theta_i \right) / n \) and with value \( \theta_i \) for buyer \( i \). With a change of variables \( z_i \equiv -x_i \) we have the results for demand competition. The (inverse) supply of the asset is given by \( p = \alpha + \beta \sum_i z_i \) (with \( \tilde{\theta} > \alpha > 0, \beta > 0 \)) where \( \sum_i z_i \) is the total quantity demanded. The equilibrium demand is \( Z(s, p) = \left( E[\theta | s, p] - p \right) / (d + \lambda) = -b + \alpha s - \beta p \) (with the endogenous parameters as in Proposition 1). The marginal benefit of buying \( z_i \) units of the asset for buyer \( i \) is \( \theta_i - \lambda z_i \), where \( \theta_i \) is the value with a private component and \( \lambda z_i \) a transaction or opportunity cost, or risk aversion component.\(^{56}\) The profits of buyer \( i \) are given by \( \pi_i = (\theta_i - p) z_i - \lambda z_i^2 / 2 \). A real market example would be firms purchasing labor of unknown average productivity \( \tilde{\theta} \) because of technological uncertainty, and facing an inverse linear labor supply and quadratic adjustment costs in the labor stock.

The extensions to the supply competition model also apply here. Of particular interest is the case of supply uncertainty since it corresponds to the *noise trader model* when \( \rho = 1 \).

Suppose that noise traders have a price-elastic demand (negative supply) \( \beta^{-1} (\alpha + u - p) \), then market clearing implies that \( \beta^{-1} (\alpha + u - p) + \sum_i z_i = 0 \) and therefore \( p = \alpha + u + \beta \sum_i z_i \). It follows from Proposition 8 that increasing noise trading (\( \sigma_u^2 \)) increases \( c \), decreases \( d \) and the expected margin (of expected marginal benefit over

---

\(^{56}\) Note that the adjustment cost is exogenous while with CARA preferences, for example, it would be endogenous and would depend, in expectation, on the degree of risk aversion times the variance of \( \theta \) conditional on the information of the trader.
price) while $c$ and $d$ are non-monotone in $\sigma^2$. When supply is inelastic and random according to $u$ and $\lambda = 0$ we recover the Kyle (1989) model with $n > 2$ risk neutral informed investors. Considering noise of the form $\beta u$ we have that demand $\beta^{-1}(\alpha + \beta u - p) \rightarrow u$ as $\beta \rightarrow \infty$ and market clearing is given by $u + \sum z_i = 0$. It follows then that in equilibrium we have an inverse of market depth (the Kyle lambda) $\partial p/\partial u = (nc)^{-1}$, exactly as in Kyle (1989), decreasing in $\sigma^2$, increasing in $\sigma^2$, and non-monotone in $\sigma^2$.\textsuperscript{57}

6. Applications

In this section we provide several applications of the model: electricity markets, strategic trade policy, pollution damages, revenue management, and financial markets. We consider supply schedule competition examples first followed by demand schedule competition cases and look at fit to our model, link to the results, and empirical evidence.

6.1 Supply schedule competition

Wholesale electricity markets

The day-ahead or spot market, with separate auctions for each delivery period (half-hourly or hourly), and the balancing market, that secures that demand and supply match at each point in time, fit our model since they are typically organized as uniform price multiunit auctions. Supplies are discrete while smooth in our model.\textsuperscript{58} The continuous (linear in particular) supply approach has been widely used and empirically implemented.

\textsuperscript{57} See Claim S.2 in Section S.5 of the Supplement (Vives (2011a)) and exercise 5.1 in Vives (2008) for the Kyle model with risk neutral investors.

\textsuperscript{58} Holmberg et al. (2008) show that if prices are selected from a discrete grid, where (realistically) the number of price levels is small in comparison to the number of quantity levels, then the step functions converge to continuous supply functions as the number of steps increase. This justifies the approximation of step-functions with smooth supply functions. The modeling of the auction with discrete supplies leads to existence problems of equilibrium in pure strategies (see von der Fehr and Harbord (1993)).
in electricity markets.\textsuperscript{59} In our base model the random residual demand a firm faces is due to cost uncertainty. Demand uncertainty is a relevant factor in the wholesale market and then our extension (Section 5.2) applies.\textsuperscript{60} Private cost information related to plant availability will be relevant when there is a day-ahead market organized as a pool where firms submit hourly or daily supply schedules. The residual demand faced by a firm will be random (even with predictable demand) since the supply of other firms depends on plant availability, which is random. The firm may have privileged information because of technical issues, transport problems, hydro availability in the reservoirs, and the terms of supply contracts for energy inputs or imports.\textsuperscript{61} Furthermore, in an emission rights system, future rights allocations may depend on current emissions and firms may have different private estimates of such allocation. This will affect the opportunity cost of using current emission rights.

The empirical evidence points to firms bidding over marginal costs.\textsuperscript{62} The Cournot framework has been used often but tends to predict prices that are too high given realistic estimates of the demand elasticity. Our model helps understand the biases introduced when taking the Cournot modeling short-cut when firms compete in supply functions (see Section 3.5). There is also evidence of information aggregation: Mansur and White (2009) show how a centralized auction market in the Eastern US yields very important information aggregation benefits over bilateral trading in order to achieve an efficient allocation in a situation where differences in marginal costs and production are private information among firms. In our model prices are revealing of average cost conditions

\textsuperscript{59} See Green and Newbery (1992), and Green (1996, 1999). See Niu et al. (2005), Hortaçsu and Puller (2008), and Sioshansi and Oren (2007) for the Texas balancing market (ERCOT).

\textsuperscript{60} In a wholesale electricity market the demand intercept $\alpha$ is a continuous function of time (load-duration characteristic) that yields the variation of demand over the time horizon considered. At any time there is a fixed $\alpha$ and the market clears. In the British Pool up to 2001, the first liberalized wholesale market, generators had to submit a single supply schedule for the entire day. Over this period residual demand facing a firm may vary considerably due to demand uncertainty and plant outages.

\textsuperscript{61} The latter include constraints in take-or-pay contracts for gas where the marginal cost of gas is zero until the constraint—typically private information to the firm—binds, or price of transmission rights in electricity imports depending on the private arrangements for the use of a congested interconnector.

\textsuperscript{62} See e.g. Borenstein and Bushnell (1999), Borenstein et al. (2002), Green and Newbery (1992), and Wolfram (1998).
(Proposition 1) but strategic behavior on the basis of private information prevents the achievement of an efficient allocation (propositions 3 and 4).

We have seen also (Proposition 2) how increasing the noise in the private signal $\sigma^2_\epsilon$ makes the slope of supply steeper (when $\rho > 0$). This result may help explain the fact that in the Texas balancing market small firms use steeper supply functions than those predicted by theory and that those departures explain the major portion of losses in productive efficiency (Hortaçsu and Puller (2008)). Indeed, smaller firms may have signals of worse quality because of economies of scale in information gathering while residual non-contract private cost information has not been taken into account in the estimation. Consistently with our analysis the welfare losses due to the “excess steepness” of supply functions over and above standard market power may be more important than the losses due to the latter. Finally, it is worth noting that the usual restriction to upward sloping schedules in electricity markets caps the market power of sellers in the spot market.

**Other interpretations of the cost shock**

The cost shock $\theta_i$ could be related to a linear ex post pollution or emission damage which is assessed on the firm and for which the producer has some private information. The regulator can introduce a quadratic subsidy to production to eliminate the distortion originating in private information (Proposition 6) and can alleviate the distortion by disclosing the available information on the average damage $\bar{\theta}$ (Proposition 9).

The cost shock can also be interpreted as a random opportunity cost of serving the market which is related to dynamic considerations (e.g. *revenue management* on the face of

---

63 The authors explain the finding by the complexity faced by small firms in set up the bidding (and argue that to take a linear approximation to marginal costs in the Texas electricity market is reasonable).

64 The rules typically require producers to submit nondecreasing (step) function offers (although in some markets, like the Amsterdam Power Exchange, retailers may submit nonincreasing demands).
products with expiration date and costly capacity changes). The value of a unit in a shortage situation is the opportunity cost of a sale. A high opportunity cost is an indication of high value of sales in the future. In this case a firm would have a private assessment of the opportunity cost with which it would form its supply schedule. For example, if supply function competition provides a suitable reduced form for pricing for airline travel then taking into account the information aggregation role of price may help explaining pricing patterns which have proved difficult to explain with extant theoretical models (see e.g. McAfee and te Velde (2006)). For example, when airlines see prices going up they may infer, correctly, that the opportunity cost is high (i.e. that expected next period demand is high) and they reduce supply in the present period to be able to supply next period at a higher profit.

The cost shock could also be a (negative) linear subsidy in a strategic trade policy game where governments manipulate the supply function of domestic firms with tariffs and subsidies. Laussel (1992) considers a market with linear demand and constant marginal costs where firms compete in a common foreign market with the help of the domestic government imposing a quadratic export tax and a linear subsidy. If the amount of subsidy is uncertain and the domestic firm receives a private noisy signal about it, we can conclude (Proposition 2 when \( \rho > 0 \)) that increasing noise in the signals softens

---

65 Situations where the product, be it a hotel room, airline flight, generated electricity or tickets for a concert, has an expiration date and capacity is fixed well in advance and can be added only at high marginal cost.

66 Talluri and Van Ryzin (2004, p. 523) state: “A typical booking process proceeds as follows. An airline posts availability in each fare class to the reservation systems stating the availability of seats in each fare class.” This is indeed like a supply function.

67 More in general, strategic agency models where an owner provides incentives to the manager to compete in the market place have typically a reduced form which is a supply function. This is similar to the presence of adjustment costs in certain industries committing the firms to supply functions (e.g. internal incentives in management consulting). See Vickers (1985), Fershtman and Judd (1987) and Faulí-Oller and Giralt (1995).

68 The quadratic tax makes steeper the slope of the effective marginal cost schedule of a firm softening competition (this would determine the \( \lambda \) in our model) and the subsidy allows the domestic firm to capture a larger share of the profits. Grant and Quiggin (1997) study the case in which firms are competitive. Whenever supply functions are linear the authors find an equilibrium in tax-subsidy schedules with quadratic trade revenue taxes.
competition. It follows that the disclosure policy of the government towards national firms can affect also competitiveness.

6.2 Demand schedule competition

*Legacy loans auctions*

They were envisioned in the US Public-Private Investment Program (March 2009) to remove bad loans from the balance sheet of banks. Basically, the banks would nominate pools of legacy loans that meet certain criteria and that they wish to sell. Approved private investors bid for the pools of loans and receive a non-recourse loan having as collateral the same securities to be acquired. The winning bid for each pool then is either accepted or rejected by the bank. In terms of our model the marginal valuation of a bidder will depend on the collateral that it can post. Using the same securities as collateral of the non-recourse loan to finance the purchase is equivalent to providing a subsidy to bidders that decreases the slope of their marginal valuation. Our model would rationalize then the subsidy scheme of the Treasury since it would reduce the discount of the auction price (Proposition 6).

*Liquidity auctions*

In an open-market central bank operation the (often inelastic) supply of funds is met by bank’s demand bids. The marginal unit value \( \theta_i \) of funds for bank \( i \) is idiosyncratic, with a common component \( \tilde{\theta} \) related to the interest rate/price in the secondary interbank market, and is assessed imperfectly by the bank (for example, due to uncertainty about future liquidity needs). Banks’ marginal valuations are positively correlated, declining with \( \lambda \) reflecting the structure of their pool of collateral (see e.g. Ewerhart et al. (2010)). A bidder bank prefers to post illiquid collateral in exchange for funds, and with an increased allotment the bidder must offer more liquid types of collateral which have a higher opportunity cost.

69 The supply function model can also account for Paulson’s reverse auction plan to extract toxic assets from the banks: it would serve a price discovery purpose but the Treasury would be subject to overpricing. Any information the Treasury has on average valuations should be released (Proposition 9). (See Section 5 in Vives (2010)). Ausubel and Cramton (2009) propose to combine the forward auction with a reverse auction as in the Paulson plan.
The model illustrates the impact in the auction discount and in the efficiency of liquidity distribution of changes in key parameters such as those happening in a crisis situation. The more severe the information problem (a larger $\rho$ or $\sigma^2_e/\sigma^2_\theta$) or the more costly to part with more liquid collateral (higher $\lambda$), the steeper demand functions are, the larger the equilibrium margin, and the inefficiency in funds allocation (and the equilibrium may break down in the inelastic case; propositions 1, 2, 3 and 4, and Section 5.1). The effects have been corroborated empirically. The central bank can try to reduce the inefficiency in the distribution of liquidity, which can be substantial when $\rho$ and/or $\sigma^2_e/\sigma^2_\theta$ are large, accepting lower quality collateral from the banks in the repo auctions. This is equivalent to provide a quadratic subsidy to the banks that lowers effectively $\lambda$. The amount of the subsidy will be increasing with $\rho$ and $\sigma^2_e/\sigma^2_\theta$ and decreasing with $n$ (Proposition 6 and Section 5.1). Central banks in the crisis have enlarged acceptable collateral and some of them the qualifying participants in the auctions.

**Treasury auctions**

The sources of private information in this context are different expectations about the future resale value of securities $\tilde{\theta}$ (for instance, bidders with different forecast of inflation with securities denominated in nominal terms), and private values arising out of different liquidity needs due to idiosyncratic shocks. There is evidence of prices in Treasury auctions featuring a discount from secondary market prices, increasing with the noise in the signal of the bidders (as in Proposition 2), as well as aggregating information.

---

70 The comparative static predictions of the auction discount with respect to the expected secondary market value $E[\tilde{\theta}|\tilde{z}]$ are consistent with documented features of the ECB euro auctions. (See Ewerhart et al. (2010), and Vives (2010). Note however that the ECB auctions in the period studied are discriminatory while our model is uniform price.) Cassola et al. (2009) show that in ECB auctions after the subprime crisis in August 2007 marginal valuations for funds of banks increased, the aggregate bid curve was steeper with increased bid shading. In our model the level effect on valuations would be represented by an increase in $\tilde{\theta}$.

71 For example, the Federal Reserve established the TAF auction facility (with a single-price format) to broaden the range of counterparties and the range of collateral in relation to regular open market operations.

Underpricing is thought to be a serious problem in uniform price auctions. There are also worries that in the financial crisis of 2007-8 margins and profits of (Wall Street) dealers may have grown dramatically at the expense of the Treasury and the Fed. Underpricing and high expected profits are consistent with our results (propositions 2 and 4).

7. Concluding remarks

In a model with private and common value uncertainty and without noise traders we find a unique privately revealing equilibrium where traders rely on their private signals (and where the incentives to acquire information are preserved). A main result is that private information generates market power over and above the full information level. Several testable implications derive from the analysis. An increase in the correlation of cost parameters or in the noise in private signals makes supply functions steeper and increases expected price-cost margins. The average margin may be above the Cournot level and get closer to the collusive level as correlation increases, with no coordination of sellers. When demand is uncertain an increase in noise will decrease expected margins.

The results may help explain pricing patterns arising in electricity markets, revenue management, and auctions. For example, ignoring private cost information with supply function competition in electricity markets may underestimate the slope of supply; and the Treasury may overpay in reverse auctions for toxic assets which reveal their average value due to increased correlation of values for banks. The biases introduced by a Cournot model when competition is in fact in supply or demand schedules are also characterized.

73 See Cammack (1991) and Nyborg et al. (2002). Gordy (1999) argues that bidders in Treasury auctions submit demand schedules to protect against the winner’s curse and he associates larger bid dispersion with increased incidence of the winner’s curse. Bid dispersion in our model can be linked to the slope of the demand schedule, with a steeper slope with more noise in the signal.

74 See the evidence provided by Kandel et al. (1999) and by Keloharju et al. (2005). US Treasury auctions are exclusively uniform price since October 1998, and only a limited number of primary dealers can submit competitive bids; in Treasury auctions in Sweden the range of participants is from 6 to 15 (Nyborg et al. (2002)).
With regard to welfare, at the SFE sellers supply too little and too similar quantities; the efficient allocation can be obtained with price taking behavior; and typically the expected deadweight loss is increasing in the correlation of the cost parameters and in the noise of private signals. With regard to policy, price taking behavior may be induced with an optimal quadratic subsidy, and a precise enough public signal about the common value component may restore efficiency. The former explains, for example, how in a financial crisis loosening collateral requirements in central bank liquidity auctions may be part of an optimal subsidy scheme to banks, or how in the US PPIP scheme for legacy loans auctions a subsidy to bidders may be rationalized.

The model and results have already been seen robust to a number of extensions maintaining the symmetry in the model. Further work should explore the role of asymmetries in technology and information structure.\footnote{Rostek and Weretka (2010) present results in the latter vein.}
Appendix: Proofs of propositions 1, 2, 3, 4, 5, and Remark 2. Statement and proof of Claims A.1, A.2, A.3 and of Lemma A.1.

The proofs of propositions 7, 8 and 9 as well as complementary material, simulations, and the analysis of information acquisition can be found in the Supplement (Vives (2011a)).

Proof of Proposition 1: (i) Suppose that sellers other than \( i \) use the strategy \( X(s_j, p) = b - a s_j + c p \). From the market clearing equation and from the point of view of seller \( i \) (provided \( 1 + \beta(n-1)c > 0 \)) the price is informationally equivalent to \( h_i \equiv \beta b(n-1) - \alpha + (1 + \beta(n-1)c) p + \beta x_i = \beta a \sum_{j \neq i} s_j \). The pair \( (s_i, p) \) is informationally equivalent to the pair \( (s_i, h_i) \), hence \( E[\theta_i | s_i, p] = E[\theta_i | s_i, h_i] \). From our Gaussian information structure

\[
\begin{pmatrix}
\theta_i \\
\bar{s}_i \\
h_i
\end{pmatrix} \sim N
\begin{pmatrix}
\bar{\theta} \\
\bar{\sigma} \\
\beta a(n-1) \bar{\theta}
\end{pmatrix}
\begin{pmatrix}
\sigma_\theta^2 & \sigma_\theta^2 & \beta a(n-1) \rho \sigma_\theta^2 \\
\sigma_\theta^2 & \sigma_\theta^2 + \sigma_\varepsilon^2 & \beta a(n-1) \rho \sigma_\theta^2 \\
\beta a(n-1) \rho \sigma_\theta^2 & \beta a(n-1) \rho \sigma_\theta^2 & \beta^2 a^2 (n-1) \left( (\sigma_\theta^2 + \sigma_\varepsilon^2) + (n-2) \rho \sigma_\theta^2 \right)
\end{pmatrix}
\]

and the projection theorem for normal random variables we obtain

\[
E[\theta_i | s_i, h_i] = \bar{\theta} + \frac{\sigma_\theta^2 \left( \sigma_\theta^2 (1 - \rho) (1 + (n-1) \rho) + \sigma_\varepsilon^2 \right)}{\left( \sigma_\theta^2 (1 - \rho) + \sigma_\varepsilon^2 \right) \left( \sigma_\theta^2 (1 + (n-1) \rho) + \sigma_\varepsilon^2 \right)} (s_i - \bar{s}) + \frac{\sigma_\theta^2 \rho \sigma_\varepsilon^2}{\left( \sigma_\theta^2 (1 - \rho) + \sigma_\varepsilon^2 \right) \left( \sigma_\theta^2 (1 + (n-1) \rho) + \sigma_\varepsilon^2 \right)} (h_i - \beta a (n-1) \bar{s}).
\]

Using the F.O.C. \( p - E[\theta_i | s_i, p] = (d + \lambda) x_i \), \( X(s_i, p) = b - a s_i + c p \) and \( h_i = p (1 + \beta n c) - \alpha + \beta n b - \beta a s_i \), we obtain the following
\[-\sigma_i^2 \left( \sigma_i^2 + (1-\rho)\sigma_i^2 \right) \bar{\theta} + \left( \sigma_i^2 \rho (\beta_{ib} - \alpha) / \beta a \right) \left( \sigma_i^2 + (1+(n-1)\rho)\sigma_i^2 \right) \]
\[-\sigma_i^2 \left( \sigma_i^2 (1-\rho) \right) \left( 1+(n-1)\rho \right) + \sigma_i^2 \right) \left( \sigma_i^2 (1-\rho) + \sigma_i^2 \right) \]
\[-\left( \sigma_i^2 (1-\rho) + \sigma_i^2 \right) \left( \sigma_i^2 (1+(n-1)\rho) + \sigma_i^2 \right) \]
\[+ \left( 1 - \frac{\sigma_i^2 \sigma_i^2 \rho (1+ \beta nc)}{\left( \sigma_i^2 + (1-\rho)\sigma_i^2 \right) \left( \sigma_i^2 + (1+(n-1)\rho)\sigma_i^2 \right) \beta a} \right) p \]
\[= (d + \lambda) \left( b - a \sigma_i + cp \right) \]

Identifying coefficients, letting \( M \equiv \frac{\rho \sigma_i^2 n}{(1-\rho) \left( \sigma_i^2 + (1+(n-1)\rho)\sigma_i^2 \right)} \), we obtain \( a, b, c \) :
\[a = \frac{(1-\rho)\sigma_i^2}{\left( \sigma_i^2 + (1-\rho)\sigma_i^2 \right)} \left( d + \lambda \right) \]
\[b = \frac{1}{1+M} \left( \frac{\alpha}{\beta n} M - \frac{\sigma_i^2}{\sigma_i^2 + (1+(n-1)\rho)\sigma_i^2} \left( d + \lambda \right)^{-1} \bar{\theta} \right) \]

with \( d = \left( \beta^2 + (n-1)c \right)^{-1} \) and where \( c \) is given by the equation
\[c = \left( d + \lambda \right)^{-1} - \frac{M \left( 1+ \beta nc \right)}{\beta n} \]
or, equivalently, the quadratic \( g(c; M) \) with:
\[g(c; M) \equiv \lambda \beta (n-1)(1+M)c^2 + \left( \beta + \lambda \right)(1+M) + \left( n-1 \left( \frac{\lambda M}{n} - \beta \right) \right) c + \frac{\beta + \lambda}{\beta n} M - 1. \]

For \( n=1 \) there is a unique solution to the quadratic equation. For \( n \geq 2 \) and \( \lambda > 0 \) the discriminant of \( g(\cdot ; M) = 0 \) is positive, and therefore the equation has two real roots, but only the largest root
\[c = \frac{-(n-M+2Mn)\lambda - n\beta (M-n+2) + \sqrt{n^2 (M-n+2)^2 \beta^2 + 2n\lambda (M+n)^2 \beta + \lambda^2 (M+n)^2}}{2n\beta \lambda (n-1)(M+1)} \]
is compatible with the second order condition \( 2d + \lambda > 0 \).

It is easily checked also that \( g \left( -M \left( (1+M)\beta n \right)^{-1}; M \right) < 0 \) and therefore for the largest root we have \( c > -M \left( (1+M)\beta n \right)^{-1} \) because of convexity of the parabola \( g(\cdot ; M) \). It
follows also that \( c > -M\left((1+M\beta n)^{-1} \right) > -(\beta n)^{-1} \) and therefore \( 1+\beta nc > 0 \) and 
\( 1+\beta(n-1)c > 0 \) either for \( c > 0 \) or \( c < 0 \), and \( 0 < d < \beta n \). Furthermore, \( a > 0 \) since \( \rho < 1 \). (Note also that \( c < \lambda^{-1} \) when \( \rho \geq 0 \) since then \( M \geq 0 \) and \( c = (d + \lambda)^{-1} - M\left( (\beta n)^{-1} + c \right) < \lambda^{-1} \) since \( 0 < d \) and \( c + (\beta n)^{-1} > 0 \).) We show that \( c \) decreases in \( \lambda \). Direct computation shows that 
\[
\frac{\partial c_1}{\partial \lambda} = \frac{\beta n (n-M-2)^2 + \lambda (M+n)^2 + (n-M-2)\sqrt{n^2(M-n+2)^2\beta^2 + 2n\lambda (M+n)^2} + \lambda^2 (M+n)^{3/2}}{2(n-1)(M+1)\sqrt{n^2(M-n+2)^2 \beta^2 + 2n\lambda (M+n)^2} + \lambda^2 (M+n)^{3/2}} < 0
\]
because \( M+1 > 0 \) and the numerator in the fraction is positive since 
\[
\left( \beta n (n-M-2)^2 + \lambda (M+n)^2 \right)^2 - \left( (n-M-2)\sqrt{n^2(M-n+2)^2 \beta^2 + 2n\lambda (M+n)^2} + \lambda^2 (M+n)^{3/2} \right)^2 
= 4\lambda^2 (M+1)(n-1)(M+n)^2 > 0.
\]
The largest root of \( g(\cdot; M) = 0 \) decreases with \( M \) since \( \partial g/\partial M > 0 \) for \( c > -(\beta n)^{-1} \).
This is so since it can be checked that \( \partial g/\partial M \) is a convex parabola in \( c \) with largest root \( -(\beta n)^{-1} \). It follows that for \( c > -(\beta n)^{-1} \) we have that \( \partial g/\partial M > 0 \). As \( M \to \infty \) we have that \( 1+\beta nc \to 0 \) and \( c \to -(\beta n)^{-1} \) since otherwise from \( c = (d + \lambda)^{-1} - \frac{M(1+\beta nc)}{b_n} \) and \( 0 < d < \beta n \) we would have that \( c \to -\infty \) which contradicts \( 1+\beta nc > 0 \). It follows that as \( M \to \infty \), \( (1+\beta(n-1)c) \to 1/n \) or \( d \to \beta n \). As \( M \to -1 \), \( -(1+M\beta n)^{-1} \to \infty \) and therefore \( c \to \infty \) and \( d \to 0 \).
(ii) If \( \lambda = 0 \) we have that at the candidate SFE \( d = \beta n(n-M-2)/(M-3n) \) (and \( c = c_0 = -(n-M)/(n-M-2)\beta n) \)). Then \( d > 0 \) (fulfilling the S.O.C.) if and only if \( n-M-2 < 0 \). ♦
Claim A.1: Let $\sigma^2 > 0$, then the absolute value of the weight on $h_i$ in $E[\theta_i | s_i, h_i]$ increases in $|\rho|$ and $\sigma^2 / \sigma_0^2$.

Proof: The coefficient of $h_i$ in $E[\theta_i | s_i, h_i]$ equals $(d + \lambda)M/\beta n$ and is increasing in $M$ since $d$ is increasing in $M$. If follows that when $\sigma^2 > 0$ the coefficient is increasing in $\rho$ since then $M$ is increasing in $\rho$ and when $M > 0$ ($M < 0$) it is increasing (decreasing) in $\sigma^2 / \sigma_0^2$ since then $M$ is increasing (decreasing) in $\sigma^2 / \sigma_0^2$. The result follows.

Claim A.2: When $\sigma^2 / \sigma_0^2 \to \infty$ we have that $a \to 0$ and $c \to \tilde{c}$, where

$$
\tilde{c} = \beta(n - 2) - \lambda - 2\rho(\beta + \lambda)(n - 1) + \sqrt{\beta^2(n - 2^2 - 2\rho(n - 1)) + \lambda^2 + 2n\beta\lambda} \over 2\beta\lambda(n - 1)((n - 1)\rho + 1),
$$

with $\tilde{c} = c'$ for $\rho = 0$ and $c' > \tilde{c}$ ($c' < \tilde{c}$) for $\rho > 0$ ($\rho < 0$) where

$$
c' = \beta(n - 2) - \lambda + \sqrt{\beta^2(n - 2^2) + 2n\beta\lambda + \lambda^2} \over 2\beta\lambda(n - 1).$$

Furthermore, $\tilde{c} \to - (\beta n)^{-1}$ for $\rho \to 1$, $\tilde{c} \to \infty$ for $\rho \to -(n - 1)^{-1}$, and $\tilde{c}$ is decreasing in $\rho$.

Proof: Note that if $\rho \sigma^2 = 0$ we have that $M = 0$ and $c$ is given by the positive root of

$$
\left(\beta^{-1} + (n - 1)c\right)^{-1} + \lambda = c^{-1}.
$$

This is $c' = \beta(n - 2) - \lambda + \sqrt{\beta^2(n - 2^2) + 2n\beta\lambda + \lambda^2} \over 2\beta\lambda(n - 1)$. When

$\sigma^2 / \sigma_0^2 \to \infty$ it is immediate that $a \to 0$, and since $M \to \rho n/(1 - \rho)$ from the expression for $c_2$ we obtain $c \to \tilde{c}$, where with $\tilde{c} = c'$ for $\rho = 0$, $\tilde{c} = - (\beta n)^{-1}$ for $\rho = 1$, $\tilde{c} \to \infty$ for $\rho \to -(n - 1)^{-1}$, and $\tilde{c}$ decreasing in $\rho$.

Claim A.3: If $n - M - 2 \geq 0$ then as $\lambda \to 0$, $c \to \infty$ and if $n - M - 2 < 0$ then as $\lambda \to 0$, $c \to c_0$.
Proof: It is immediate from the expression for \( c_2 \) that when \( n - M - 2 \geq 0 \), \( \lim_{\lambda \to 0} c_2 = \infty \), and when \( n - M - 2 < 0 \) using l'Hôpital rule we find that

\[
\lim_{\lambda \to 0} c_2 = \lim_{\lambda \to 0} \frac{(\lambda + n\beta)(M + n)^2}{n^2(M-n+2)^2\beta^2 + 2n\lambda(M+n)^2\beta + \lambda^2(M+n)^2} - \frac{(n-M+2\beta\lambda)}{2n\beta(n-1)(M+1)} = -\frac{n-M}{(n-M-2)\beta n}.
\]

\[\blacktriangleleft\]

Proof of Proposition 2:

(i) The slope \( c \) decreases with \( M \) which, when \( \rho \sigma_\epsilon^2 > 0 \), increases with \( \rho \) and \( \sigma_\theta^2/\sigma_\epsilon^2 \).\(^{76}\) When \( \rho < 0 \), \( M \) decreases with \( \sigma_\epsilon^2/\sigma_\theta^2 \) and \( c \) increases in \( \sigma_\epsilon^2/\sigma_\theta^2 \). As \( \rho \) ranges from \(-1\) to \(1\), \( M \) ranges from \(-1\) to \(\infty\) and \( c \) ranges from \(\infty\) to \(-1/\beta n\), and \( d \) ranges from \(0\) to \(\beta n\) (from Proposition 1).

(ii) In equilibrium, \( p = E[\hat{\theta}|\hat{s}] + (d + \lambda)\hat{x} \) and from the demand function we obtain \( \hat{x} = (\alpha - E[\hat{\theta}|\hat{s}])/(\beta n + \lambda + d) \). If sellers share the signals \( s = (s_1, \ldots, s_n) \) and maximize joint profits they solve:

\[
\max_{(x_i)_{i=1}^n} E\left[\sum_{i=1}^n \pi_i | s\right]
\]

where \( E\left[\sum_{i=1}^n \pi_i | s\right] = \sum_{i=1}^n p x_i - \sum_{i=1}^n E[\theta_i | s] x_i - \frac{\lambda}{2} \sum_{i=1}^n x_i^2 \), with (sufficient) F.O.C.

\[
\alpha - 2\beta \sum_{i=1}^n x_i - E[\theta_i | s] - \lambda x_i = 0, \quad i = 1, \ldots, n.
\]

Adding up across sellers we obtain

\[
\frac{p - E_n[M\pi_n]}{p} = \frac{1}{\eta}.
\]

It is immediate that the average collusive output is \( \bar{x}^n = (\alpha - E[\hat{\theta}|\hat{s}] / (2\beta n + \lambda) \). At the SFE as \( \rho \to 1 \) (when \( \sigma_\epsilon^2 > 0 \)) we have that \( d \to \beta n \), and for given \( \hat{s} \) the aggregate

\[^{76}\) Note that the equilibrium depends only on the ratio \( \sigma_\epsilon^2/\sigma_\theta^2 \).
interim Lerner index converges to the collusive level \( \eta^{-1} \) and \( \bar{x} \to \bar{x}^* \). Similarly, as \( \rho \to -(n-1)^{-1} \) we have that \( d \to 0 \) and \( \bar{x} \to \bar{x}^* \equiv \left( \alpha - E[\tilde{\theta}|\tilde{s}] \right)/(\beta n + \lambda) \), the average competitive output.

(iii) We have that \( \bar{p} - E[MC_n] = dE[\tilde{x}] = d \left( \alpha - \tilde{\theta} \right) \left( \beta n + \lambda + d \right) \), which is increasing in \( d \), and therefore with \( \rho \) or \( \sigma_{\varepsilon}^2/\sigma_\theta^2 \) (when \( \rho > 0 \)). From Claim A.4 below \( d \) decreases with \( n \) when \( \rho \geq 0 \) and \( c > 0 \). The same happens with the expected price since \( \bar{p} = \tilde{\theta} + (d + \lambda) \left( \alpha - \tilde{\theta} \right) \left( \beta n + \lambda + d \right) \) is increasing in \( d \). The opposite results for \( \sigma_{\varepsilon}^2/\sigma_\theta^2 \) when \( \rho < 0 \) since then \( d \) is decreasing in \( \sigma_{\varepsilon}^2/\sigma_\theta^2 \).

(iv) We have that \( \text{var}[\tilde{x}] = (\beta n + \lambda + d)^2 \text{var} \left[ E \left[ \tilde{\theta}|\tilde{s} \right] \right] \) where \( E \left[ \tilde{\theta}|\tilde{s} \right] = \zeta \tilde{s} + (1 - \zeta) \tilde{\theta} \) and \( \zeta \equiv \text{var}[\tilde{\theta}]/\left( \text{var} \left[ \tilde{\theta} \right] + \sigma_{\varepsilon}^2 n^{-1} \right) \). It follows that \( \text{var} \left[ E \left[ \tilde{\theta}|\tilde{s} \right] \right] = \zeta^2 \text{var} \left[ \tilde{s} \right] = \zeta \text{var}[\tilde{\theta}] = \frac{\left( (1 + (n-1)\rho) \sigma_{\theta}^2 \right)^2}{\left( (1 + (n-1)\rho) \sigma_{\theta}^2 + \sigma_{\varepsilon}^2 \right) n} \) increases in \( \rho \) (since \( \rho > -(n-1)^{-1} \)) and \( \sigma_\theta^2 \), and decreases in \( \sigma_{\varepsilon}^2 \). We conclude that price volatility \( \text{var}[p] = (\beta n)^2 \text{var}[\tilde{x}] \) decreases with \( \sigma_{\varepsilon}^2 \) and increases with \( \sigma_\theta^2 \) when \( \rho > 0 \) (since \( d \) increases with \( \sigma_{\varepsilon}^2/\sigma_\theta^2 \) when \( \rho > 0 \)).

Claim A.4: \( d \) decreases with \( n \) when \( \rho \geq 0 \) and \( c > 0 \).

Proof: At the equilibrium \( \tilde{\theta} = \tilde{\theta} \) and it is possible to check that \( \frac{\partial c}{\partial n} = -\frac{\partial g}{\partial c} \) =

\[
\frac{(1 + \lambda d^{-1})(1 + cn\beta)\left( M^{-1}(1 - \rho) - \rho \right) M}{n(\rho(n-1) + 1)} + c(M^{-1}(\beta + 2\lambda) - (\beta - c\lambda)(1 + M(2 - n^{-1}))) \]
\[
- \frac{2c\lambda\beta(n-1)(\beta + M + 1) + (\beta + \beta\lambda)(M + 1) + (M^{-1}\beta - \beta)(n-1))}{2c\lambda\beta(n-1)(\beta + M + 1) + (\beta + \beta\lambda)(M + 1) + (M^{-1}\beta - \beta)(n-1)}.
\]

It follows that \( \frac{\partial}{\partial n}\left( c(n-1) \right) = (n-1)\frac{\partial c}{\partial n} + c =

\[
\frac{\beta(1 + \lambda d^{-1})(M^{-1}(\rho - M^{-1}(1 - \rho))(M^{-1} + c\beta(M + 1)) + c\beta(M + 1))}{(\rho(n-1) + 1)\beta(\partial g/\partial c)} > 0
\]
when \( c > 0 \) since \( \partial g / \partial c > 0 \), and when \( \rho \geq 0 \) we have that \( M \geq 0 \) and \( \rho - Mn^{-1}(1 - \rho) \geq 0 \). Note that for \( \rho = 0 \), \( d = d^f \) and therefore \( d^f (c^f) \) is decreasing (increasing) in \( n \). ♦

Proof of Proposition 3:
(i) Considering a Taylor series expansion of \( TS \) (stopping at the second term due to the quadratic nature of the payoff) around the efficient allocation, which maximizes \( E[TS] \),\(^{77}\) it follows that
\[
E[DWL] = n \left( \beta n (\tilde{x}(t;d) - \tilde{x}(t;0))^2 + \lambda n^{-1} \sum_i (x_i(t;d) - x_i(t;0))^2 \right) / 2.
\]

To supply an average quantity \( \tilde{x} \) the market solves the program
\[
\min_{(\xi, \tilde{\xi})} \left\{ E \left[ \sum_{i=1}^n \hat{C}(x_i, \theta) | t \right] \right\} \text{ s.t. } n^{-1} \sum_{i=1}^n x_i = \tilde{x}
\]
yielding \( \hat{x}_i = \tilde{x} + (\tilde{t} - t_i)(d + \lambda)^{-1}, \ i = 1, \ldots, n \).\(^{78}\) If \( \tilde{x} \) and \( \tilde{x}^0 \) are supplied in a cost-minimizing way then \( x_i - \tilde{x} = x_i^0 - \tilde{x}^0 = (\tilde{t} - t_i)\lambda^{-1} \) and \( x_i - x_i^0 = \tilde{x} - \tilde{x}^0 \). Since
\[
\tilde{x}(t;d) = (\alpha - \tilde{t})/(\beta n + \lambda + d),
\]
it follows that aggregate inefficiency is given by
\[
n \left( \beta n + \lambda \right) (\tilde{x}(t;d) - \tilde{x}(t;0))^2 \right) / 2 = \left( \beta n + \lambda \right) \left( (\beta n + \lambda)^{-1} - (\beta n + \lambda + d)^{-1}\right)^2 (\alpha - \tilde{t})^2 / 2
\]
and it is increasing in \( d \). The residual in the deadweight loss is due to distributive inefficiency. Letting \( u_i = x_i(t;d) - \tilde{x}(t;d), \ u_i^0 = x_i(t;0) - \tilde{x}(t;0), \) and \( \tilde{\sigma}_i^2 = n^{-1} \sum_i (t_i - \tilde{t})^2 \), and noting that \( u_i = (\tilde{t} - t_i)/(\lambda + d) \) and \( u_i^0 = (\tilde{t} - t_i)/\lambda \), distributive inefficiency is given by
\[
\lambda \sum_i (u_i - u_i^0)^2 / 2 = n\lambda \left( \lambda^{-1} - (\lambda + d)^{-1}\right)^2 \tilde{\sigma}_i^2 / 2,
\]
and it is increasing in \( d \).

\(^{77}\) The result holds true, in fact, comparing an efficient allocation with any another allocation which is based on weakly coarser information. See Lemma 1 in Vives (2002).

\(^{78}\) Note that \( E[\theta | k] = n^{-1} \sum_{i=1}^n E[\theta_i | k] = n^{-1} \sum_{i=1}^n t_i \equiv \tilde{t} \).
(ii) Let \( \pi_i(t; d) = E(\pi_i | t) \), where \( \pi_i \) are profits at the SFE, and 
\( \tilde{\pi}(t; d) \equiv n^{-1} \sum_i \pi_i(t; d) \), then it can be shown that in equilibrium 
\( \pi_i(t; d) = \left( d + \frac{\lambda}{2} \right) (x_i(t; d))^2 \) where \( x_i(t; d) = \bar{x}(t; d) + (\bar{t} - t_i)(d + \lambda)^{-1} \). It follows that 
\[
\tilde{\pi}(t; d) = \left( d + \frac{\lambda}{2} \right) \left( \bar{x}(t; d) \right)^2 + \frac{\sigma_i^2}{(\lambda + d)^2}
\]
and 
\[
\frac{\partial \tilde{\pi}(t; d)}{\partial d} = \frac{\beta n - d}{\beta n + d + \lambda} \left( \bar{x}(t; d) \right)^2 - \frac{d}{(\lambda + d)^3} \frac{\sigma_i^2}{}.\]
This implies that \( \frac{\partial \tilde{\pi}}{\partial d} > 0 \) for \( d \) small and \( \frac{\partial \tilde{\pi}}{\partial d} < 0 \) for \( d \) close to \( \beta n \). ♦

Proof of Proposition 4: Let \( x_i = x_i(t; d) \) and \( x_i^0 = x_i(t; 0) \). From the proof of Proposition 3(i) we obtain
\[
E[DWL] = E[E[DWL | t]] = n \left( \beta n E \left[ (\bar{x} - \bar{x}^0)^2 \right] + \lambda E \left[ (x_i - x_i^0)^2 \right] \right)/2,
\]
and the corresponding decomposition
\[
E[DWL] = n \left( (\beta n + \lambda) E \left[ (\bar{x} - \bar{x}^0)^2 \right] + \lambda E \left[ (u_i - u_i^0)^2 \right] \right)/2,
\]
with \( u_i \equiv x_i - \bar{x} \) and \( u_i^0 \equiv x_i^0 - \bar{x} \), where the first term corresponds to aggregate inefficiency and the second to distributive inefficiency.

(i) We have that \( E \left[ (\bar{x} - \bar{x}^0)^2 \right] = \left( (\beta n + \lambda)^{-1} - (\beta n + \lambda + d)^{-1} \right)^2 E \left[ (\alpha - \bar{\alpha})^2 \right] \). We know that increases in \( \rho \) or \( 1/\sigma_x^2 \) increase the variance of the prediction \( \bar{t} = E[\bar{t} | \tilde{s}] \) (see proof of Proposition 2(iv)), and therefore \( E \left[ (\alpha - \bar{\alpha})^2 \right] \) increases in \( \rho \) and decreases in \( \sigma_x^2 \). Since \( d \) increases in \( \rho \) (for \( \sigma_x^2 > 0 \)) we can conclude that aggregate inefficiency increases with \( \rho \). We have also that \( d \) increases in \( \sigma_x^2 \) if \( \rho > 0 \) and it can be checked that aggregate inefficiency may be non-monotonic with respect to \( \sigma_x^2 \). If \( \rho \leq 0 \) then \( d \) is weakly decreasing in \( \sigma_x^2 \) and aggregate inefficiency decreases in \( \sigma_x^2 \).
(ii) We have that $E \left[ (u_i - u_i^0)^2 \right] = \left( \lambda^{-1} - (\lambda + d)^{-1} \right)^2 E \left[ (t_i - \bar{t})^2 \right]$. I claim that $E \left[ (t_i - \bar{t})^2 \right]$ is decreasing in $\rho$ and $\sigma_e^2$ when $\rho < 1$. Noting that

$$t_i = E[\theta | s_i, \bar{s}] = \bar{\theta} + \frac{(1-\rho)\sigma_{\theta}^2}{\sigma_{\theta}^2(1-\rho) + \sigma_e^2}(s_i - \bar{\theta}) + \frac{\sigma_{\theta}^2\rho n}{((n-1)\rho + 1)\sigma_{\theta}^2 + \sigma_e^2}(\bar{\theta})$$

and $\tilde{t} = \bar{\theta} + \frac{(1-\rho)\sigma_{\theta}^2}{\sigma_{\theta}^2(1-\rho) + \sigma_e^2}(\bar{s} - \bar{\theta}) + \frac{\sigma_{\theta}^2\rho n}{((n-1)\rho + 1)\sigma_{\theta}^2 + \sigma_e^2}(\bar{\theta})$, we obtain

$$t_i - \tilde{t} = \frac{(1-\rho)\sigma_{\theta}^2}{\sigma_{\theta}^2(1-\rho) + \sigma_e^2}(s_i - \bar{s})$$

and $E \left[ (t_i - \tilde{t})^2 \right] = \text{var}[t_i - \tilde{t}] = \left( \frac{(1-\rho)\sigma_{\theta}^2}{\sigma_{\theta}^2(1-\rho) + \sigma_e^2} \right)^2 \text{var}[s_i - \bar{s}]$. Since $\text{var}[s_i - \bar{s}] = (\sigma_{\theta}^2(1-\rho) + \sigma_e^2)(n-1)n^{-1}$, we conclude that $E \left[ (t_i - \tilde{t})^2 \right] = \frac{(1-\rho)^2(n-1)\sigma_{\theta}^4}{n(\sigma_{\theta}^2(1-\rho) + \sigma_e^2)}$, which is decreasing in $\rho$ and $\sigma_e^2$ (and increasing in $n$) when $\rho < 1$. The effect of $\rho$ ($\sigma_e^2$) on $E \left[ (t_i - \tilde{t})^2 \right]$ is particularly strong when $\sigma_e^2$ ($\rho$) is small (close to 0), while $d$ increases in $\rho$ and $\sigma_e^2$ (when $\rho > 0$). The total effect can go either way. If $\sigma_e^2 = 0$ then $d$ is independent of $\rho$ and distributive inefficiency decreases in $\rho$. If $\rho \leq 0$ then $d$ is weakly decreasing in $\sigma_e^2$ and distributive inefficiency decreases in $\sigma_e^2$.

(iii) From the proof of Proposition 3 (ii) we have that

$$E \left[ \pi_i(t;d) \right] = E \left[ \tilde{\pi}(t;d) \right] = \left( d + \frac{\lambda}{2} \right) \left( E \left[ (\alpha - \bar{\lambda})^2 \right] + \frac{E \left[ (t_i - \tilde{t})^2 \right]}{(\beta n + \lambda + d)^2} \right).$$

When $\rho \to 1$ and $\sigma_e^2 > 0$ we know (from Proposition 2) that $d \to \beta n$ (and $\tilde{x}(t,d) \to \tilde{x}^m = \tilde{x}(t,\beta n)$). Furthermore, as $\rho \to 1$, $E \left[ (t_i - \tilde{t})^2 \right] \to 0$, $x_i - \tilde{x} = (\tilde{t} - t_i)(d + \lambda)^{-1}$ tends in mean square to 0 and productive inefficiency at the SFE vanishes. (Note that $x_i^m = \tilde{x}^m + (\tilde{t} - t_i)\lambda^{-1}$ and therefore as $\rho \to 1$ $x_i \to x_i^m = \tilde{x}^m$ in mean square.) It follows that $E \left[ \pi_i(t;d) \right]$ converge to the collusive level as $\rho \to 1$:
\[
\left(\beta n + \frac{\lambda}{2}\right)E\left[\left(\alpha - \tilde{\eta}\right)^2\right].
\]
When \(\rho \to - (n-1)^{-1}\) and \(\sigma^2 > 0\), then \(d \to 0\)
\((\tilde{x}(t,d) \to \tilde{x}(t,0)\) and \(x_i(t,d) \to x_i(t,0)\), and \(E[\pi_i(t;d)]\) converge to the
competitive level \(E[\pi_i(t;d = 0)]\). (Note also that \(\text{var}[]\to 0\) as \(\rho \to - (n-1)^{-1}\).)
When \(\sigma^2 = 0\), \(E[\pi_i]\) is linear and decreasing in \(\rho\) since then
\[\text{sgn}\left\{\partial E[\pi_i]/\partial \rho\right\} = \text{sgn}\left\{\frac{((\beta n + \lambda + d)^2 - (\lambda + d)^2)}{\sigma^2}\right\} < 0.\]
If \(\rho \leq 0\) then \(d\) is weakly
decreasing in \(\sigma^2\) and an increase in \(\sigma^2\) leads to a decrease in both \(E[\left(\alpha - \tilde{\eta}\right)^2]\) and
\(E[\left(t_i - \tilde{\eta}\right)^2]\). ♦

Proof of Remark 2: Similarly as in the proof of Proposition 4 we obtain
\[ETS^\nu - ETS^f = n\left((\beta n + \lambda)E\left[\left(\tilde{x}^f - \tilde{x}^\nu\right)^2\right] + \lambda E\left[\left(u^f_i - u^\nu_i\right)^2\right]\right)/2,\]
and since neither \(\rho\) nor \(\sigma^2\) affect \(d^f\) it follows that both aggregate and distributive
inefficiency decrease in \(\sigma^2\) and therefore the deadweight loss is decreasing in \(\sigma^2\), and
aggregate (distributive) inefficiency increases (decreases) in \(\rho\).

Proof of Proposition 5:
(i) A price-taking SFE is Bayesian equilibrium where price-taking is imposed. The
equilibrium strategy of seller \(i\) is of the form \(X^{PT}(s_i, p) = b^{PT} - a^{PT} s_i + c^{PT} p\) and it arises
out of the maximization of expected profits taking prices as given but using the
information contained in the price:
\[\max_{x_i}\left\{(p - E[\theta_i | s_i, p])x_i - \frac{\lambda}{2}x_i^2\right\}.\]
Following the same procedure as in the proof of Proposition 1 but with \(d = 0\) we obtain
the equilibrium. In the equilibrium we have that \(c^{PT} = \left(\lambda^{-1} - \left(\beta n\right)^{-1} M\right)/(M + 1),\)
\(1 + \beta \nu c^{PT} > 0\), \(a^{PT} > 0\), \(p = \left(1 + \beta \nu c^{PT}\right)^{-1}\left(\alpha - \beta \nu b^{PT} + \beta n^{PT} \tilde{s}\right)\) and \(p\) reveals \(\tilde{s}\).
(ii) For a given \( \kappa < \lambda \) and induced slope of marginal cost \( \lambda - \kappa > 0 \), we know that \( \partial c / \partial \kappa > 0 \) since \( c \) is decreasing in \( \lambda \) (Proposition 1). It follows that \( d(\kappa) \) is decreasing in \( \kappa \) up to \( \kappa = \lambda \). However, the fixed-point equation \( d(\kappa) = \kappa \) need not have a solution unless allowing for negative slopes of effective marginal costs. We have that

\[
c^\text{PT}(\lambda) = \left(\lambda^{-1} - (\beta n)^{-1} M\right)(M+1)^{-1}
\]

goes from \( +\infty \) to \( -M \left(1 + M) \beta n \right)^{-1} \) as \( \lambda \) moves in the range \((0, +\infty)\). Therefore given that \( c^\text{PT}(\hat{\lambda}) > c(\hat{\lambda}) \) and that both are decreasing in \( \lambda \), for any \( \lambda > 0 \) there is always a \( \kappa > 0 \) such that \( c(\lambda - \kappa) = c^\text{PT}(\lambda) \) provided that the range of \( c(\lambda) \) is the same as \( c^\text{PT}(\lambda) \) (see Figure 4a). This is so if and only if \( n - M - 2 \geq 0 \). In this case as \( \lambda \to 0, \ c(\lambda) \to \infty \). If \( n - M - 2 < 0 \) then as \( \lambda \to 0, \ c(\lambda) \to c_0 \), where \( c_0 = -(n-M)/(n-M-2) \beta n \). (See Claim A3.) Then only if \( \lambda \) is such that \( c^\text{PT}(\lambda) \leq c_0 \) we can find the desired \( \kappa > 0 \) with \( \lambda - \kappa \geq 0 \). Otherwise we need to induce \( \lambda - \kappa < 0 \) to obtain \( c(\lambda - \kappa) = c^\text{PT}(\lambda) \). This is feasible since if \( n - M - 2 < 0 \) and \( \hat{\lambda} < \lambda < 0 \), \( \hat{\lambda} = -\frac{n \beta}{M + n} \left(M + n - 2\sqrt{M+1(n-1)}\right) \), then there are two linear SFE, with slopes of supply \( c_1 > c_2 \lim_{\lambda \to 0^-} c_1 = +\infty, \lim_{\lambda \to 0^-} c_2 = c_0 \), and

\[
\lim_{\lambda \to \hat{\lambda}} c_1 = \lim_{\lambda \to \hat{\lambda}} c_2 = \hat{c} = \frac{(n-M)\sqrt{(M+1)(n-1) + (M+n)}}{\beta n \sqrt{(M+1)(n-1) - n + M + 2}}, \quad \partial c_1 / \partial \lambda > 0 \quad \text{and} \quad \partial c_2 / \partial \lambda < 0 \quad \text{(see Lemma A.1 below).}
\]

Therefore for \( n - M - 2 < 0 \) and \( c^\text{PT}(\hat{\lambda}) > c_0 \) we can always find a \( \kappa > 0 \) such that \( c(\lambda - \kappa) = c^\text{PT}(\lambda) \): If \( c^\text{PT}(\hat{\lambda}) > \hat{c} \) take \( \kappa > 0 \) such that \( c_1(\lambda - \kappa) = c^\text{PT}(\lambda) \); if \( c^\text{PT}(\hat{\lambda}) \leq \hat{c} \) then take \( \kappa > 0 \) such that \( c_2(\lambda - \kappa) = c^\text{PT}(\lambda) \). (See Figure 4b.) The optimal subsidy is increasing in \( M \) and \( \lambda \) (since \( c^\text{PT} \) decreases with \( M \) and \( \lambda \)). This means in particular that \( \kappa^* \) increases with \( \rho \) and increases (decreases) with \( \sigma^2_e \) when \( \rho > 0 \) (\( \rho < 0 \)).\*
Lemma A.1. If \(n - M - 2 < 0\) and \(\hat{\lambda} < \lambda < 0\), where
\[
\hat{\lambda} = -n \frac{\beta}{M+n} \left(M + n - 2\sqrt{(M+1)(n-1)}\right),
\]
then there are two roots of \(g(c; M) = 0\) that fulfill the S.O.C., with \(c_1 > c_2\), \(\partial c_1 / \partial \lambda > 0\), \(\partial c_2 / \partial \lambda < 0\), \(\lim_{\lambda \to 0} c_1 = +\infty\), \(\lim_{\lambda \to 0} c_2 = c_0\), and
\[
\lim_{\lambda \to \hat{\lambda}} c_1 = \lim_{\lambda \to \hat{\lambda}} c_2 = \frac{(n-M)\sqrt{(M+1)(n-1)} + (M+n)}{\beta n \sqrt{(M+1)(n-1)(M+2-n)}}.
\]
When \(\lambda = \hat{\lambda}\) there is only one root and it fulfills the S.O.C.

Proof: Let \(c_1\) denote the smallest root of \(g(c; M) = 0\) when \(\lambda > 0\). We have that
\[
c_2 - c_1 = \frac{\sqrt{n} (M-n+2) \beta^2 + 2n\lambda (M+n)^2 \beta + \lambda^2 (M+n)^2}{n\beta \lambda (n-1)(M+1)}.
\]
When \(\hat{\lambda} < \lambda < 0\) the discriminant is also positive, and if \(n - M - 2 > 0\) it can be checked that no root fulfills the S.O.C. while if \(n - M - 2 < 0\) both roots do (and the largest solution is \(c_1\)). When \(\lambda = \hat{\lambda}\) there is only one root and it fulfills the S.O.C. If \(n - M - 2 = 0\) then \(\hat{\lambda} = 0\). If \(\hat{\lambda} < \lambda < 0\) and \(n - M - 2 < 0\) then from the expression for \(c_1\), \(\lim_{\lambda \to 0} c_1 = \infty\) (and indeed \(\lim_{\lambda \to 0} c_2 = c_0\)). Direct computation yields that
\[
\lim_{\lambda \to \hat{\lambda}} c_1 = \lim_{\lambda \to \hat{\lambda}} c_2 = \frac{(n-M)\sqrt{(M+1)(n-1)} + (M+n)}{\beta n \sqrt{(M+1)(n-1)(M+2-n)}}.
\]
Furthermore,
\[
\frac{\partial c_1}{\partial \lambda} = \frac{\beta n (n-M-2)^2 + \lambda (M+n)^2 - (n-M-2)\sqrt{n^2(M-n+2)\beta^2 + 2n\lambda (M+n)^2 \beta + \lambda^2 (M+n)^2}}{2(n-1)(M+1)\sqrt{n^2(M-n+2)\beta^2 + 2n\lambda (M+n)^2 \beta + \lambda^2 (M+n)^2 \lambda^2}} < 0
\]
whenever \((n-M-2) < 0\) since
\[
\beta n (n-M-2)^2 + \lambda (M+n)^2 \geq \beta n (n-M-2)^2 + \hat{\lambda} (M+n)^2 = \frac{\beta n (n-M-2)^2 + \lambda (M+n)^2}{2(n-1)(n+1)+\sqrt{(M+1)(n-1)(M+n)}} \geq 0
\]
as \(-2(M+1)(n-1)) + \sqrt{(M+1)(n-1)(M+n)} \geq 0\).

References


